



# Adjunctive treatment with oral dexamethasone in non-ICU patients hospitalised with community-acquired pneumonia: a randomised clinical trial

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**Adjunctive treatment with oral dexamethasone in adults hospitalised with community-acquired pneumonia (CAP) reduced length of stay and ICU admission rate. However, it remains unclear for which CAP subgroup the risk–benefit ratio is optimal.** <https://bit.ly/35tXfPX>

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## Abstract

**Background** Adjunctive intravenous corticosteroid treatment has been shown to reduce length of stay (LOS) in adults hospitalised with community-acquired pneumonia (CAP). We aimed to assess the effect of oral dexamethasone on LOS and whether this effect is disease severity dependent.

**Methods** In this multicentre, stratified randomised, double-blind, placebo-controlled trial, immunocompetent adults with CAP were randomly assigned (1:1 ratio) to receive oral dexamethasone (6 mg once daily) or placebo for 4 days in four teaching hospitals in the Netherlands. Randomisation (blocks of four) was stratified by CAP severity (pneumonia severity index class I–III and IV–V). The primary outcome was LOS.

**Results** Between December 2012 and November 2018, 401 patients were randomised to receive dexamethasone (n=203) or placebo (n=198). Median LOS was shorter in the dexamethasone group (4.5 days, 95% CI 4.0–5.0 days) than in the placebo group (5.0 days, 95% CI 4.6–5.4 days; p=0.033). Within both CAP severity subgroups, differences in LOS between treatment groups were not statistically significant. The secondary ICU admission rate was lower in the dexamethasone arm (5 (3%) versus 14 (7%); p=0.030); 30-day mortality did not differ between groups. In the dexamethasone group the rate of hospital readmission tended to be higher (20 (10%) versus 9 (5%); p=0.051) and hyperglycaemia (14 (7%) versus 1 (1%); p=0.001) was more prevalent.

**Conclusion** Oral dexamethasone reduced LOS and ICU admission rate in adults hospitalised with CAP. It remains unclear for which patients the risk–benefit ratio is optimal.

## Introduction

Despite advances in antibiotic treatment and the availability of preventative measures such as vaccines, the burden of community-acquired pneumonia (CAP) remains high [1]. Therefore, nonantibiotic adjunctive therapies that modify the host response to microorganisms remain of interest [2].

Excessive release of cytokines in response to invading pathogens is thought to contribute to high mortality and morbidity in patients with CAP [3]. Corticosteroids can inhibit inflammation by downregulating this cytokine response [4]. Through this mechanism, adjunctive treatment with corticosteroids might improve clinical outcomes.

Several randomised controlled trials (RCTs) show that adjunctive corticosteroid treatment reduces hospital length of stay (LOS) [5]. However, most RCTs have studied intravenous corticosteroid treatment. Dexamethasone administered *i.v.* during the first 4 days of hospitalisation has been shown to reduce LOS by 1 day [6]. Oral administration of dexamethasone has several advantages over *i.v.* administration. It does not hamper an early *i.v.*-to-oral switch of antibiotics, causes patients less discomfort and carries no risk of phlebitis. Furthermore, a bioequivalence study showed that oral dexamethasone is feasible from a pharmacokinetics perspective [7]. Thus, we opted to investigate the effect of oral dexamethasone in patients with CAP.

Moreover, it is still debated which patients benefit most from corticosteroid treatment [8]. A recent individual patient data meta-analysis (IPDMA) suggested a greater effect of corticosteroids in patients with severe CAP, defined by a high pneumonia severity index (PSI) score [5]. So far, no RCT has prospectively investigated the effects of corticosteroids in pre-specified subgroups based on CAP severity.

The primary objective of this study was to investigate the effect of a short course of oral dexamethasone compared with placebo on LOS and to assess whether this effect depends on disease severity.

## Materials and methods

### Study design and patients

This multicentre, stratified randomised, double-blind, placebo-controlled trial was conducted in four nonacademic teaching hospitals in the Netherlands. Patients presenting with CAP were screened and enrolled within 24 h of emergency department presentation. Inclusion criteria were age  $\geq 18$  years and the presence of new opacities on chest radiography, and two of the following signs and symptoms: cough, production of sputum, temperature  $>38.0^{\circ}\text{C}$  or  $<36.0^{\circ}\text{C}$ , abnormalities at auscultation consistent with pneumonia, C-reactive protein (CRP)  $>15\text{ mg}\cdot\text{L}^{-1}$ , white blood cell count  $>10\times 10^9$  or  $<4\times 10^9\text{ cells}\cdot\text{L}^{-1}$ , or  $>10\%$  of bands in leukocyte differentiation. The following patients were excluded from study participation: patients with a congenital or acquired immunodeficiency, patients treated with chemotherapy  $<6$  weeks prior to emergency department presentation, patients receiving corticosteroids or other immunosuppressive medication 6 weeks prior to emergency department presentation, patients requiring direct admission to the intensive care unit (ICU) at hospital presentation, patients with a known tropical worm infection, pregnant or breastfeeding females and patients with an intolerance for dexamethasone. Patients opting for palliative care, who did not receive active treatment for pneumonia, were also not eligible for study participation. All other patients with limitations in treatment (*e.g.* those who did not wish to be resuscitated, or did not want to be admitted to the ICU if necessary, or those who did not wish to be intubated) but who did seek active treatment for the pneumonia were eligible for study participation. Written informed consent was provided by all patients. This study was approved by the Medical Ethics Committee at St Antonius Hospital (Nieuwegein, The Netherlands) and is registered at ClinicalTrials.gov with identifier number NCT01743755.

Eligible patients were randomly allocated (1:1 ratio) to receive either 6 mg oral dexamethasone or placebo once a day for 4 days. A previous pharmacokinetics study showed that 6 mg dexamethasone orally equals the exposure of 5 mg dexamethasone phosphate (=4 mg dexamethasone) *i.v.*, as studied in the Ovidius trial [6, 7]. Randomisation was performed in blocks of four using PASW Statistics software version 18.0.03 (IBM, Armonk, NY, USA). Patients were stratified by enrolling centre and by CAP severity (nonsevere CAP and severe CAP). Nonsevere CAP was defined as PSI class I–III and severe CAP was defined as PSI class IV–V [9]. Randomisation was set up to ensure that in each CAP severity subgroup, 50% of patients received dexamethasone and 50% of patients received placebo. After randomisation, patients were assigned a medication kit number using a central computer-assisted allocation system. Corresponding coded medication kits containing four tablets of 6 mg dexamethasone or placebo were available at the emergency department of each of the participating hospitals. Patients, treating physicians and investigators were masked to treatment allocation.

### Methods

Patients in the dexamethasone group received 6 mg oral dexamethasone (Tiofarma, Oud-Beijerland, The Netherlands) once a day for 4 days and patients in the placebo group received one placebo tablet (Tiofarma) once a day for 4 days. Study treatment was initiated within 24 h of emergency department

presentation. Baseline blood samples for blood chemistry testing and haematology were obtained before initiation of study treatment in the emergency department as part of standard care. Measurements included CRP, electrolytes, glucose, renal function and a complete blood count. All patients received antibiotics prior to starting study medication. Decisions regarding antibiotic type, route of administration and treatment duration were made by the treating physician, and were based on Dutch national guidelines [10, 11]. Microbiological testing included sputum cultures, blood cultures, PCR assays for respiratory viruses and atypical pathogens, and urinary antigen tests for the detection of *Legionella pneumophila* serogroup 1 and *Streptococcus pneumoniae*. The decision to transfer a patient to the ICU or to discharge a patient was made by the treating physician. The general rule for discharge in all hospitals was that patients were clinically stable (improvement of shortness of breath, consistent decrease in CRP concentrations, absence of hyperthermia or hypothermia, adequate oral intake and adequate gastrointestinal absorption) and in well enough condition to leave the hospital. Baseline characteristics included medical history and variables necessary to calculate the PSI score [9].

### Analysis

The primary outcome was LOS measured in 0.5 days. LOS was calculated from time of emergency department presentation to the day of discharge, day of death or day of ICU admission (study medication was stopped after ICU admission because patients are regularly treated with corticosteroids in the ICU). If the patient was admitted to the emergency department before 12:00, the day of presentation was counted as 1 day. If the patient was admitted to the emergency department after 12:00, the day of presentation was counted as 0.5 days. The discharge date was defined as the date that a patient was medically ready for discharge (hereby excluding waiting time for admission to a nursing home). Time of discharge was set at 12:00 for all patients as patients are generally discharged late morning or early afternoon depending on ward logistics. Secondary outcomes were admission to the ICU after initial admission to the general ward and all-cause mortality within 30 days of hospital admission.

Sample size estimation was based on our hypothesis that dexamethasone could reduce the median LOS in all patients with CAP by 1 day and reduce the median LOS in patients with severe CAP by 2 days. With sample data pseudo-randomly generated from available data from our previous trial [6] and assuming that 50% of patients have severe CAP, it was simulated that 300 patients were needed in each arm to provide >80% power maintaining a type I error rate of 0.05 (two-sided).

The primary analysis was a Kaplan–Meier analysis of time to discharge. The Kaplan–Meier method was used to estimate the median LOS with 95% confidence interval for each treatment group and to assess the difference in LOS between treatment groups by analysing time to discharge. Patients who died, who were transferred to a different hospital or who were admitted to the ICU after study enrolment were censored to show that time of reporting was cut off before the event of interest for the primary analysis (*i.e.* hospital discharge) occurred. Because the intervention was a short course of oral dexamethasone, a Gehan–Breslow–Wilcoxon test was used for the Kaplan–Meier method as this test emphasises early differences [12]. Furthermore, we performed an extra sensitivity analysis in which patients who were admitted to the ICU after study enrolment were included in the time to discharge analysis.

To adhere to CONSORT (Consolidated Standards of Reporting Trials) guidelines on reporting results of randomised clinical trials we also calculated the unadjusted hazard ratio (HR) for discharge with 95% confidence interval using a Cox proportional hazards regression [13]. Differences in secondary outcomes between treatment groups were analysed with a Chi-squared test and risk ratios were calculated; a two-tailed p-value <0.05 was deemed significant. Statistical analyses were performed using SPSS version 24.0 (IBM). The primary analysis was performed according to the intention-to-treat principle after which the analysis was repeated in the per-protocol population. Patients who missed one or more doses of study medication while admitted to the general ward, whose diagnosis was altered, with exclusion criteria unknown at the time of study entry, or who were discharged on the day of study entry were excluded from the per-protocol analysis. The following predefined subgroup analyses were performed: 1) CAP severity (nonsevere CAP *versus* severe CAP), 2) initial CRP level at emergency department presentation (above median *versus* below median) and 3) *S. pneumoniae* urinary antigen test result.

We added a sensitivity analyses to explore the effect of dexamethasone on hospital utilisation. The difference in hospital utilisation between treatment groups was assessed using a 30-day hospital-free approach (equivalent to the mechanical ventilator-free days approach). Hospital-free days (HFDs) were calculated by adding the number of days a patient was hospitalised during readmission (if a readmission occurred within 30 days of initial hospital admission) to the duration of initial hospital stay (including ICU admission) and subtracting this number from 30 days. If a patient died in hospital within 30 days of

admission, the number of HFDs was 0. If a patient was not discharged within 30 days of admission, the number of HFDs was also 0. Because the effect of dexamethasone is primarily through shortened LOS rather than mortality, a Mann–Whitney U-test was used to compare HFDs between groups [14].

Categorical variables are shown as number (percentage). Continuous variables are presented as median (interquartile range (IQR)) or mean with standard deviation for variables with a nonparametric or parametric distribution, respectively.

Interim analyses to monitor the frequency of serious side-effects related to either dexamethasone or placebo were pre-planned at 200, 400 and 500 patients. The analyses and the review of the results were performed by an external independent Data Safety and Monitoring Board.

## Results

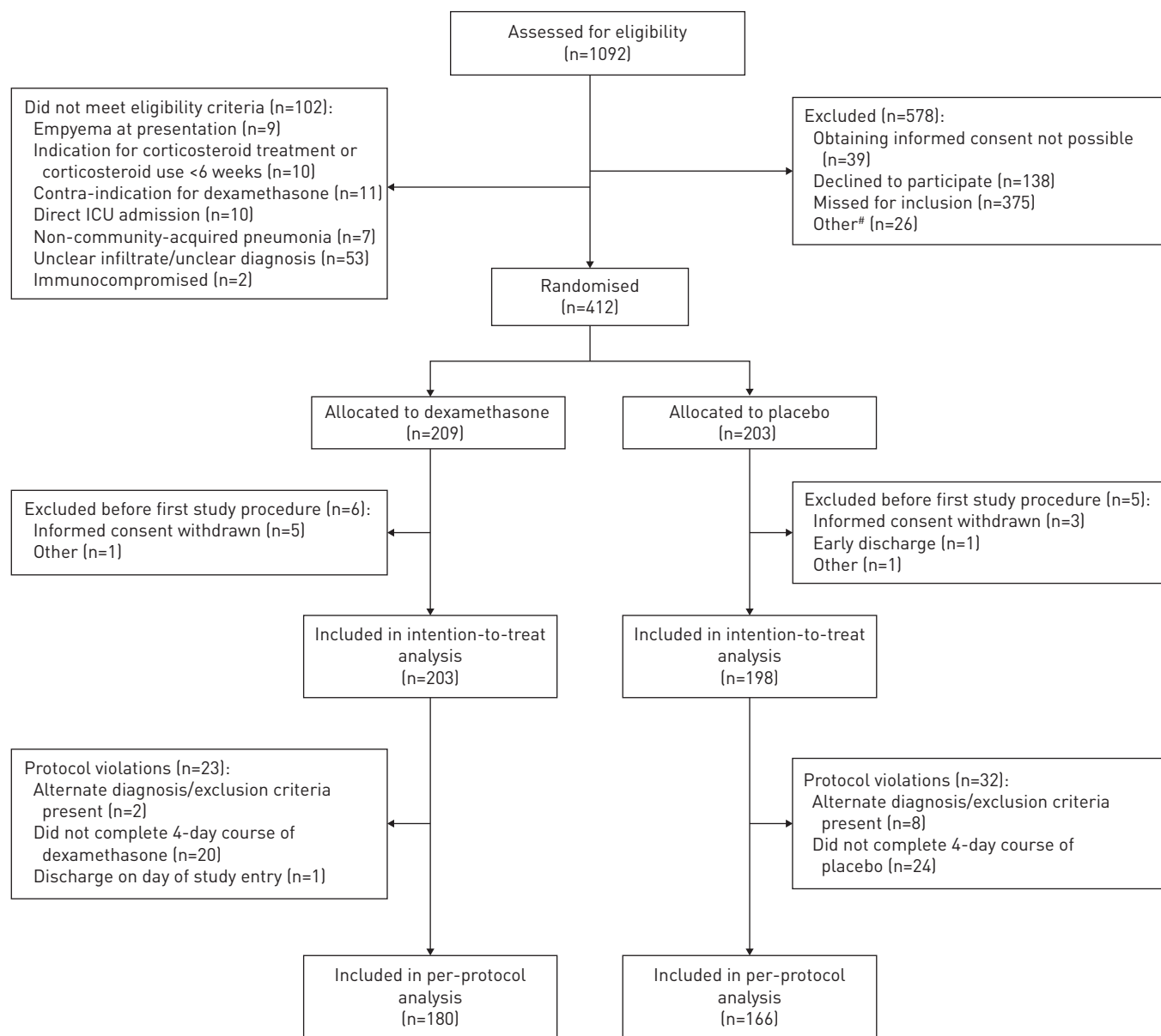
From 23 December 2012 to 28 November 2018, 1092 patients were screened for eligibility. For one hospital, screening logs were not available. 412 patients were randomly allocated to receive either dexamethasone or placebo; 11 patients were excluded post-randomisation (figure 1). The study was prematurely terminated after the second interim analysis due to a slower inclusion rate than anticipated combined with a shorter LOS than used in our sample size calculation. Therefore, we did not expect a different outcome for LOS at 600 patients. Furthermore, for 30-day mortality we anticipated a 50% lower mortality rate in patients with severe CAP in the dexamethasone group compared with the placebo group (7.5% versus 15% based on results of an earlier trial) [6]. Because there was no difference in 30-day mortality between treatment groups at 400 patients and the 30-day mortality was already lower than anticipated, we also did not expect a different outcome for 30-day mortality at 600 patients. The independent Data Safety and Monitoring Board found no grounds for early termination based on safety concerns.

There was no difference in baseline characteristics between the intervention and placebo groups (table 1). The mean $\pm$ SD PSI score calculated for all patients was 81 $\pm$ 29. The severe CAP subgroup consisted of 156 (39%) patients. There was no difference in the distribution of causative organisms and initial antibiotic treatment between treatment groups (supplementary tables E1 and E2).

In the intention-to-treat population, Kaplan–Meier analysis showed that median LOS was 0.5 days shorter in the dexamethasone group (4.5 (95% CI 4.0–5.0) days) than in the placebo group (5.0 (95% CI 4.6–5.4) days) (table 2). Kaplan–Meier analysis of time to discharge showed a significant difference between treatment groups ( $p=0.033$ ) (figure 2). Although nonstatistically significant, in the nonsevere CAP subgroup LOS was 1.0 day shorter in the dexamethasone group compared with the placebo group (table 2 and figure 3). There was no difference in LOS between treatment groups in the severe CAP subgroup (table 2 and figure 3). Results were similar in the per-protocol population (supplementary table E3). In the Kaplan–Meier analysis in which ICU patients were not censored, median LOS was 5.0 (95% CI 4.5–5.5) days in the dexamethasone group and 5.5 (95% CI 5.0–6.0) days in the placebo group ( $p=0.012$ ) (supplementary figure E1). Using Cox regression, HR for discharge was 1.14 (95% CI 0.93–1.39) for all patients, 1.19 (95% CI 0.92–1.54) in the mild pneumonia group and 1.06 (95% CI 0.76–1.48) in the severe pneumonia group.

For secondary outcomes, the secondary ICU admission rate was lower in the dexamethasone group ( $n=5$  (3%)) than in the placebo group ( $n=14$  (7%);  $p=0.030$ ). Respiratory failure was the most common reason for ICU admission (supplementary table E5). The 30-day mortality rate did not differ between both treatment groups (table 2). Causes of death are shown in supplementary table E6. The aforementioned results for the intention-to-treat population were similar in the per-protocol population (supplementary table E3). Results of predefined subgroup analyses are presented in supplementary table E4.

Adverse events are shown in table 3. The readmission rate within 30 days of study entry was higher in the dexamethasone group compared with the placebo group (20 (10%) versus 9 (5%);  $p=0.051$ ). Reasons for readmission are shown in supplementary table E7. The median (IQR) number of HFDs was 25.0 (22.0–26.0) in the dexamethasone group and 24.5 (22.5–26.5;  $p=0.061$ ) in the placebo group. Hyperglycaemia was reported by physicians in 14 (7%) patients in the dexamethasone group and one (1%) patient in the placebo group ( $p=0.001$ ). In the placebo group, one patient had a newly diagnosed myxoma and one patient was diagnosed with HIV. Both were transferred to an academic hospital. In the dexamethasone group, one patient had a perforated jejunal diverticulitis requiring surgical intervention. Abdominal complaints were present before study entry. Furthermore, in the dexamethasone group three patients had an ischaemic cerebrovascular accident and one patient developed deep venous thrombosis of the right leg.



**FIGURE 1** Study profile. No patient was lost to follow-up before reaching the primary end-point. ICU: intensive care unit. #: e.g. transferred to another hospital or patient opting for palliative care.

## Discussion

In the primary analysis of this trial, we observed a reduction in median LOS of 0.5 days in patients with CAP treated with oral dexamethasone compared with controls.

This finding supports our hypothesis that dexamethasone reduces LOS in patients with CAP. However, a 0.5-day reduction is lower than the hypothesised 1-day reduction. It is also lower than reported by BRIEL *et al.* [5] who also found a 1-day reduction of LOS in their IPDMA of six trials. The median LOS in our study was shorter compared with all trials included in the IPDMA by BRIEL *et al.* [5], which may explain the difference in absolute reduction in LOS. Still, the relative reduction in LOS was 10% in our trial compared with 12.5% found by BRIEL *et al.* [5]. Thus, the relative effect of dexamethasone on LOS in our study was similar. The difference in overall LOS could be explained by the fact that most studies in the IPDMA used *i.v.* study medication; this may have hampered early *i.v.*-to-oral antibiotic switch and consequently an earlier discharge. Furthermore, there were fewer patients with severe CAP in our trial compared with the two trials in the IPDMA with similar inclusion criteria (39% versus 47% and 49%) [6, 15].

TABLE 1 Baseline characteristics of enrolled patients

	All patients		PSI I-III		PSI IV-V	
	Placebo	Dexamethasone	Placebo	Dexamethasone	Placebo	Dexamethasone
<b>Patients</b>	198	203	119	126	79	77
<b>Male</b>	120 (61)	116 (57)	58 (49)	63 (50)	62 (79)	53 (69)
<b>Age years</b>	67 (54–76)	68 (57–76)	61 (44–69)	61 (50–70)	77 (68–83)	76 (69–83)
<b>Ethnicity</b>						
Caucasian	186 (94)	197 (97)	111 (93)	122 (97)	75 (95)	75 (97)
Other	11 (6)	6 (3)	7 (6)	4 (3)	4 (5)	2 (3)
<b>Elderly home resident</b>	1 (1)	6 (3)	0 (0)	2 (2)	1 (1)	4 (5)
<b>Current smoker</b>	45 (23)	53 (26)	26 (22)	39 (31)	19 (24)	14 (18)
<b>Antibiotic treatment prior to admission</b>	57 (29)	56 (28)	40 (34)	35 (28)	17 (22)	21 (27)
<b>Comorbidities</b>						
Neoplastic disease	6 (3)	8 (4)	1 (1)	0 (0)	5 (6)	8 (10)
Liver disease	2 (1)	2 (1)	1 (1)	1 (1)	1 (1)	1 (1)
Congestive heart failure	17 (9)	20 (10)	4 (3)	4 (3)	13 (17)	16 (21)
Renal disease	27 (14)	32 (16)	6 (5)	7 (6)	21 (27)	25 (33)
Diabetes mellitus	47 (24)	41 (20)	14 (12)	22 (18)	33 (42)	19 (25)
COPD	35 (18)	40 (20)	20 (17)	22 (18)	15 (19)	18 (23)
<b>Physical examination findings</b>						
Temperature °C	38.3±1.2	38.4±1.1	38.3±1.1	38.4±0.9	38.3±1.3	38.4±1.3
Systolic blood pressure mmHg	128±22	130±22	127±20	131±18	121 (112–147)	130 (104–148)
Heart rate beats·min <sup>-1</sup>	98 (87–110)	99 (87–111)	98 (90–110)	100 (90–111)	98±20	98±23
Respiratory rate breaths·min <sup>-1</sup>	20 (18–25)	20 (16–25)	21±5	20±5	23±7	23±7
Blood oxygen saturation %	93.6±4.1	93.7±4.2	94.6±3.7	94.1±4.4	92.2±4.2	93.0±3.6
Altered mental status	14 (7)	13 (6)	0 (0)	1 (1)	14 (18)	12 (16)
<b>Inflammatory parameters</b>						
CRP mg·L <sup>-1</sup>	198 (82–309)	211 (86–330)	190 (84–291)	249 (131–336)	203 (61–323)	153 (41–314)
WBC count ×10 <sup>9</sup> L <sup>-1</sup>	13.0 (9.7–17.5)	13.7 (10.1–18.2)	13.0 (9.6–17.6)	14.0 (10.3–19.0)	13.0 (9.7–17.1)	13.1 (9.4–17.5)
<b>PSI</b>	82±29	81±29	69 (52–76)	65 (52–76)	106 (97–115)	106 (97–120)
<b>PSI class</b>						
1	25 (13)	27 (13)	25 (21)	27 (21)		
2	40 (20)	55 (27)	40 (34)	55 (44)		
3	54 (27)	44 (22)	54 (45)	44 (35)		
4	70 (35)	64 (32)			70 (89)	64 (83)
5	9 (5)	13 (6)			9 (11)	13 (17)

Data are presented as n, n (%), median (interquartile range) or mean±sd. PSI: pneumonia severity index; COPD: chronic obstructive pulmonary disease; CRP: C-reactive protein; WBC: white blood cell.

The Cox regression analysis did not show a statistically significant difference in time to discharge between treatment groups. This analysis was included to adhere to CONSORT guidelines on reporting clinical trial results. However, the Cox regression requires the assumption of proportional hazards. Because we investigated a short course of dexamethasone and most patients were discharged during the first 5 days of hospital admission, the assumption of proportional hazards is not met.

This is the first study to show a reduction in the rate of secondary ICU admissions in patients with CAP receiving corticosteroids. However, as respiratory failure was the main reason for ICU admission (n=14 (74%)), this finding is in line with the meta-analysis by STERN *et al.* [16] who showed a lower risk of new respiratory failure in patients receiving corticosteroids. In line with the IPDMA by BRIEL *et al.* [5], we did not observe a beneficial effect of corticosteroids on 30-day mortality. STERN *et al.* [16] did show a beneficial effect of corticosteroids on mortality. However, in that meta-analysis, small studies with an unclear allocation concealment were mainly responsible for that finding [17–19].

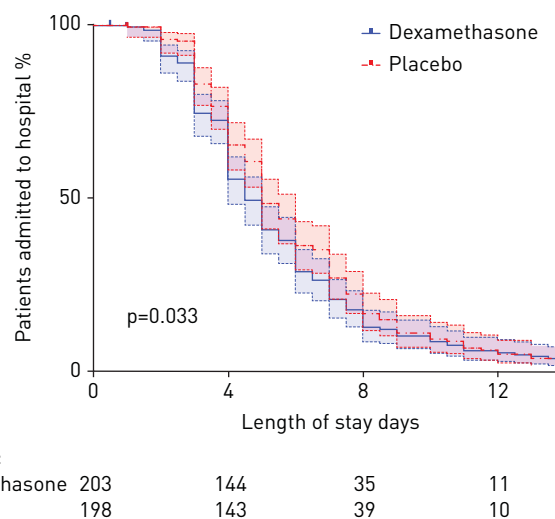
Contrary to our hypothesis, we did not observe a beneficial effect of dexamethasone in patients with severe CAP. The beneficial effects of dexamethasone seemed greater in the nonsevere CAP subgroup. In the latter group, no patients receiving dexamethasone were admitted to the ICU and the median LOS was 1.0 day shorter in patients receiving dexamethasone compared with those receiving placebo (although not statistically significant). It is difficult to draw conclusions due to the relatively small number of patients in

TABLE 2 Overview of primary and secondary end-points for the intention-to-treat population

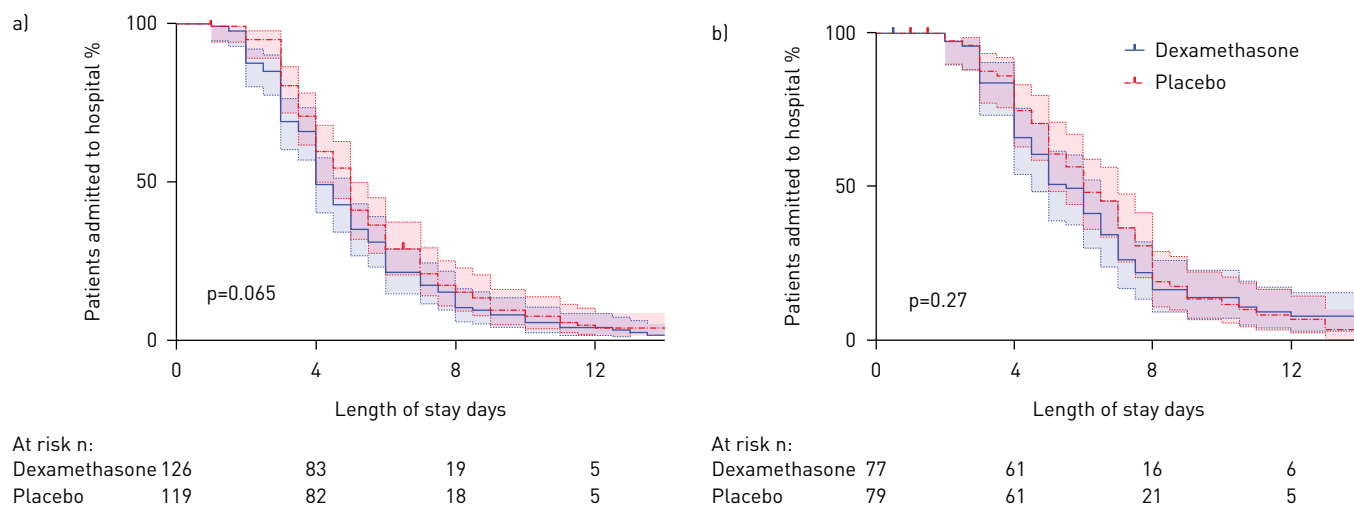
	Dexamethasone <sup>#</sup>	Placebo <sup>¶</sup>	Risk ratio (95% CI)	p-value
<b>Patients</b>	203	198		
<b>Hospital LOS days</b>				
All patients	4.5 (4.0–5.0)	5.0 (4.6–5.4)		0.033 <sup>+</sup>
PSI class I–III	4.0 (3.6–4.4)	5.0 (4.5–5.5)		0.065 <sup>+</sup>
PSI class IV–V	5.5 (4.6–6.4)	6.0 (5.1–6.9)		0.27 <sup>+</sup>
<b>Secondary ICU admission</b>				
All patients	5 (3)	14 (7)	0.35 (0.13–0.95)	0.030 <sup>§</sup>
PSI class I–III	0 (0)	6 (5)		0.011 <sup>§</sup>
PSI class IV–V	5 (7)	8 (10)	0.64 (0.22–1.87)	0.41 <sup>§</sup>
<b>30-day mortality</b>				
All patients	4 (2)	7 (4)	0.56 (0.17–1.87)	0.34 <sup>§</sup>
PSI class I–III	1 (1)	2 (2)	0.47 (0.04–5.14)	0.53 <sup>§</sup>
PSI class IV–V	3 (4)	5 (6)	0.62 (0.15–2.49)	0.49 <sup>§</sup>

Data are presented as n, median (95% CI) or n (%), unless otherwise stated. LOS: length of stay; PSI: pneumonia severity index; ICU: intensive care unit. <sup>#</sup>: PSI class I–III n=126, PSI class IV–V n=77; <sup>¶</sup>: PSI class I–III n=119, PSI class IV–V n=79; <sup>+</sup>: Gehan–Breslow–Wilcoxon test; <sup>§</sup>: Chi-squared test.

each subgroup. However, it is still interesting to explore this counterintuitive finding. It could be related to the fact that we used the PSI score to define severe CAP. The PSI score is a good predictor of mortality, yet the PSI score does not necessarily correspond with the level of inflammation. The PSI score is mainly influenced by age and the presence of comorbidities. We therefore performed an additional explorative analysis using the CURB-65 (confusion, urea  $>7$  mmol·L<sup>-1</sup>, respiratory rate  $\geq 30$  breaths·min<sup>-1</sup>, blood pressure  $<90$  mmHg (systolic) or  $\leq 60$  mmHg (diastolic), age  $\geq 65$  years) score. The CURB-65 score is based on clinical parameters; it does not include comorbidities and is less influenced by age than the PSI score. Indeed, we found the largest LOS reduction in patients aged  $<65$  years with high CURB-65 scores ( $\geq 2$  points) (supplementary figure E2). Furthermore, in our predefined subgroup analysis dexamethasone reduced LOS and the rate of secondary ICU admission in patients with a CRP above the median. We did not find this effect in patients with a CRP below the median. Two *post hoc* analyses of RCTs investigating corticosteroids in CAP have also noted that patients with a high level of inflammation benefitted most from corticosteroids. REMMELTS *et al.* [20] previously observed that dexamethasone was most effective in patients with a high level of pro-inflammatory cytokines combined with discrepantly low cortisol levels. URWYLER



**FIGURE 2** Kaplan–Meier analysis of the effect of dexamethasone on hospital length of stay in all enrolled patients. Patients who were admitted to the intensive care unit and/or died in hospital (n=21) and patients who were transferred to another hospital (n=2) were censored on the day of admission to the unit, day of death or day of transfer. The shading represents the confidence bands.



**FIGURE 3** Kaplan–Meier analysis of the effect of dexamethasone on hospital length of stay stratified according to community-acquired pneumonia severity: pneumonia severity index a) I–III and b) IV–V. Patients who died, were admitted to the intensive care unit or were transferred to a different hospital were censored on the day of death, day of admission to the unit or day of transfer. The shading represents the confidence bands.

*et al.* [21] found that only a high level of pro-inflammatory cytokines predicted a positive response to steroids. Consequently, a prediction score based solely on the level of inflammation is of interest as it might aid in identifying the subgroup of patients that would benefit most from dexamethasone.

Regarding safety, the rate of patients readmitted within 30 days of admission was twice as high in the dexamethasone group compared with the placebo group (5% versus 10%; number needed to harm n=20). However, this difference did not reach statistical significance. The rate of hyperglycaemia was higher in the dexamethasone group, which is in line with the pharmacology of corticosteroids and with an earlier trial [6].

Our study has several strengths. First, it is the second largest multicentre trial assessing the effects of corticosteroids in patients with CAP and it is the first trial to use stratified randomisation to assess the effects of corticosteroids within subgroups based on CAP severity. Second, a short course of oral dexamethasone has several advantages over longer courses of *i.v.* administered corticosteroids.

There were several limitations to this study. First, the results cannot be generalised to all patients with CAP. Patients admitted directly to an ICU (*i.e.* the most critically ill patients) were excluded. Second, the trial was prematurely terminated due to slower inclusion rates than anticipated. The results of the interim review of the study’s data at 400 patients showed a shorter LOS compared with our sample size calculation and therefore

	Dexamethasone	Placebo	Risk ratio (95% CI)	p-value <sup>+</sup>
<b>Subjects</b>	203	198		
<b>Adverse event</b>				
Readmission <sup>#</sup>	20 (10)	9 (5)	2.09 (0.98–4.47)	0.051
Empyema	3 (2)	5 (3)	0.59 (0.14–2.42)	0.45
Hyperglycaemia	14 (7)	1 (1)	13.7 (1.81–103)	0.001
Neuropsychiatric complaints ( <i>e.g.</i> delirium, agitation)	10 (5)	7 (4)	1.39 (0.54–3.59)	0.49
Cardiac events ( <i>e.g.</i> arrhythmia, congestive heart failure, myocardial infarction)	9 (4) <sup>¶</sup>	4 (2)	2.19 (0.69–7.01)	0.17

Data are presented as n or n (%), unless otherwise stated. <sup>#</sup>: n=201 patients analysed in the dexamethasone group and n=189 patients analysed in the placebo group (excluding missing (n=2) and patients who died in hospital (n=9)); <sup>¶</sup>: one patient suffered myocardial infarction and was admitted to the cardiac ward, and one patient was admitted to the cardiac ward after discharge due to ongoing angina pectoris and fatigue; <sup>+</sup>: Chi-squared test.



we do not expect a different outcome for LOS at 600 patients. Furthermore, because 30-day mortality was lower than anticipated and because there was no difference in 30-day mortality between treatment groups at 400 patients, we would not expect different findings if the planned 600 patients would have been included. Last, the number of patients reported to have hyperglycaemia is substantially lower than described by Briel *et al.* [5]. We cannot exclude the possibility of underreporting as the presence of hyperglycaemia was based on voluntarily reporting by research physicians instead of a structured assessment. Glucose was measured on day 4, a time when many patients were already discharged. In hindsight, this might limit an all-inclusive benefit–risk assessment. However, the relative risk was similar to other studies.

The benefits of dexamethasone should be weighed against the risks. A 10% reduction in LOS and reduction in ICU admissions seems to be a considerable benefit for patients. However, this should be weighed against a possible rise in readmissions. The sensitivity analysis using HFDs showed a small (statistically nonsignificant) difference between treatment groups in favour of dexamethasone. It seems that corticosteroid treatment does not benefit all patients with CAP. Therefore, it is important to identify subgroups of patients who benefit most and/or suffer least from corticosteroid treatment. High levels of inflammatory biomarkers such as cytokines, procalcitonin, pro-adrenomedullin and a high neutrophil/lymphocyte ratio have been associated with unfavourable outcomes in CAP [21–23]. In other studies, only measurement of inflammation based on cytokine levels has been shown to predict response to corticosteroids. In the present study we found that dexamethasone had a greater effect in patients with a high CRP. Future research is necessary to determine how CRP and other inflammatory biomarkers can predict response to corticosteroids, preferably using readily available biochemical tests that provide fast results.

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This study is registered at ClinicalTrials.gov with identifier number NCT01743755. Individual participant data that underlie the results reported in this article after de-identification, a dictionary defining each field in the dataset and the study protocol will be made available after approval of the proposal by the Santeon-CAP Study Group. Requests should be directed to [w.bos@antoniuziekenhuis.nl](mailto:w.bos@antoniuziekenhuis.nl) (on behalf of the Santeon-CAP Study Group) along with an analysis proposal; data requestors will need to sign a data access agreement.

Author contributions: W.J.W. Bos, S.M.C. Spoorenberg, J.C. Grutters and E.M.W. van de Garde designed the study and wrote the study protocol. S.M.T. Vestjens, S.M.C. Spoorenberg and E. Wittermans were responsible for recruitment and follow-up of the participants and study coordination. W.J.W. Bos, J.C. Grutters, F.W.J.M. Smeenk, W.L. Blok and R. Janssen were the principal investigators in the participating hospitals. G.P. Voorn was responsible for microbiological data. E.M.W. van de Garde supervised the packaging and labelling of study medication, and contributed to the data analysis. E. Wittermans analysed the data and wrote the first draft of the manuscript. W.J. W. Bos, J.C. Grutters, G.T. Rijkers, E.M.W. van de Garde, G.P. Voorn, S.M.T. Vestjens, S.M.C. Spoorenberg, F.W.J.M. Smeenk, W.L. Blok and R. Janssen critically revised the manuscript. All authors had access to the data and contributed substantially to the submitted manuscript.

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