# Circulating Levels of Insulin-like Growth Factor 1 and Insulinlike Growth Factor Binding Protein 3 Associate With Risk of Colorectal Cancer Based on Serologic and Mendelian Randomization Analyses 

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## IGF1 Levels Associate With Risk of Colorectal Cancer

# Circulating IGF1 levels: 


> Actual
> Genetically predicted
Mendelian randomization

BACKGROUND \& AIMS: Human studies examining associations between circulating levels of insulin-like growth factor 1 (IGF1) and insulin-like growth factor binding protein 3 (IGFBP3) and colorectal cancer risk have reported inconsistent results. We conducted complementary serologic and Mendelian randomization (MR) analyses to determine whether alterations in circulating levels of IGF1 or IGFBP3 are associated with colorectal cancer development. METHODS: Serum levels of IGF1 were measured in blood samples collected from 397,380 participants from the UK Biobank, from 2006 through 2010. Incident cancer cases and cancer cases recorded first in death certificates were identified through linkage to national cancer and death registries. Complete follow-up was available through March 31, 2016. For the MR analyses, we identified genetic variants associated with circulating levels of IGF1 and IGFBP3. The association of these genetic variants with colorectal cancer was examined with 2 -sample MR methods using genomewide association study consortia data ( 52,865 cases with colorectal cancer and 46,287 individuals without [controls]) RESULTS: After a median follow-up period of 7.1 years, 2665 cases of colorectal cancer were recorded. In a multivariableadjusted model, circulating level of IGF1 associated with colorectal cancer risk (hazard ratio per 1 standard deviation increment of IGF1, 1.11; 95\% confidence interval [CI] 1.051.17). Similar associations were found by sex, follow-up time, and tumor subsite. In the MR analyses, a 1 standard deviation increment in IGF1 level, predicted based on genetic factors, was associated with a higher risk of colorectal cancer risk (odds ratio 1.08; 95\% CI 1.03-1.12; $P=3.3 \times 10^{-4}$ ). Level of IGFBP3, predicted based on genetic factors, was associated with colorectal cancer risk (odds ratio per 1 standard deviation increment, 1.12; 95\% CI 1.06-1.18; $P=4.2 \times 10^{-5}$ ). Colorectal cancer risk was associated with only 1 variant in the IGFBP3 gene region (rs11977526), which also associated with anthropometric traits and circulating level of IGF2. CONCLUSIONS: In an analysis of blood samples from almost 400,000 participants in the UK Biobank, we found an association between circulating level of IGF1 and colorectal cancer. Using genetic data from 52,865 cases with colorectal cancer and 46,287 controls, a higher level of IGF1, determined by genetic factors, was associated with colorectal cancer. Further studies are needed to determine how this signaling pathway might contribute to colorectal carcinogenesis.

Keywords: CRC; Risk Factors; Signal Transduction; GWAS.

Insulin-like growth factor-1 (IGF1) has mitogenic and anti-apoptotic effects and has been implicated in the development and progression of several cancers. ${ }^{1,2}$ The bioactivity of IGF1 is partially regulated through insulinlike growth factor binding proteins (IGFBPs-), with approximately $80 \%$ bound to IGFBP3. ${ }^{3}$ In addition to its IGF-binding properties, IGFBP3 also has been shown to exhibit direct antiproliferative and pro-apoptotic effects. ${ }^{4,5}$ Multiple epidemiological studies have investigated the associations of circulating IGF1 and IGFBP3 levels with colorectal cancer risk. ${ }^{6-14}$ However, most of

## WHAT YOU NEED TO KNOW

## BACKGROUND AND CONTEXT

Human studies examining associations between circulating levels of insulin-like growth factor 1 (IGF1) and insulin-like growth factor binding protein 3 (IGFBP3) and colorectal cancer risk have reported inconsistent results

## NEW FINDINGS

In an analysis of blood samples from almost 400,000 participants in the UK Biobank, we found an association between circulating level of IGF1 and colorectal cancer. Using genetic data on 52,865 persons with colorectal cancer and 46,287 persons without, a higher level of IGF1, determined by genetic factors, was also associated with colorectal cancer. Further studies are needed to determine how this signaling pathway might contribute to colorectal carcinogenesis.

## LIMITATIONS

This is an association study between measured and genetically determined blood levels of proteins and cancer incidence.

## IMPACT

Changes in IGF signaling might contribute to colorectal carcinogenesis, but further studies are needed to determine the mechanisms of this process.
these studies were of relatively small size ( $<500$ colorectal cancer cases) and reported inconsistent results. As such, there is currently a lack of consensus as to whether higher circulating levels of IGF1 and lower levels of IGFBP3 are risk factors for colorectal cancer.

To address this, we conducted complementary serologic and Mendelian randomization (MR) analyses to examine the role of circulating IGF1 and IGFBP3 in colorectal cancer development. We investigated how prediagnostic circulating levels of IGF1 are related to colorectal cancer risk in the UK Biobank study, a large prospective cohort. We then used a 2 -sample MR approach to obtain causal estimates of the associations by combining genetic variants associated with circulating IGF1 and IGFBP3 levels in genome-wide association studies (GWAS) and then assessing the association of these variants with colorectal cancer risk in a large consortium of 52,865 colorectal cancer cases and 46,287 controls. ${ }^{15}$

Abbreviations used in this paper: BMI, body mass index; CI , confidence interval; CRP, C-reactive protein; GWAS, genome-wide association studies; HbA1c, glycolated hemoglobin; HR, hazard ratio; ICC, intraclass correlation coefficient; IGF1, insulin-like growth factor 1; IGFBP3, insulinlike growth factor binding protein 3; MR, Mendelian randomization; MRPRESSO, Mendelian Randomization Pleiotropy RESidual Sum and Outlier; OR, odds ratio; SD, standard deviation; SHBG, sex hormone binding globulin; SNP, single nucleotide polymorphism.

## Most current article

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## Methods

## UK Biobank: Serologic Analysis

Study participants. The UK Biobank is a prospective cohort study that aims to investigate the genetic, lifestyle, and environmental causes of a range of diseases. ${ }^{16,17}$ This research has been conducted using the UK Biobank Resource under application number 25897. Between 2006 and 2010, 502,656 adults aged between 40 and 69 years $(229,182$ men and 273,474 women) were recruited. All participants were registered with the UK National Health Service and lived within approximately 25 miles ( 40 km ) of 1 of the 22 study assessment centers. The UK Biobank invited approximately 9.2 million people to participate through postal invitation with a telephone follow-up, with a response rate of $5.7 \%$. The UK Biobank has approval from the North West Multi-centre Research Ethics Committee, the National Information Governance Board for Health and Social Care in England and Wales, and the Community Health Index Advisory Group in Scotland. In addition, an independent Ethics and Governance Council was formed in 2004 to oversee the UK Biobank's continuous adherence to the Ethics and Governance Framework that was developed for the study (http://www.ukbiobank.ac.uk/ethics/). All participants provided written informed consent. During the baseline recruitment visit, participants were asked to complete a self-administered touchscreen questionnaire, which included questions on sociodemographics (including age, sex, education, and Townsend deprivation score), health and medical history, lifestyle exposures (including smoking habits, dietary intakes, and alcohol consumption), early life exposures, and medication use. At the baseline visit, participants also underwent physical measurements, including body weight, height, and waist circumference. Blood samples were collected from all participants at recruitment and from a subset of approximately 20,000 participants who re-attended the assessment center between 2012 and 2013 for a repeat assessment visit. Blood samples were labeled, centrifuged, and stored at $-80^{\circ} \mathrm{C}$.

Exclusions before the onset of analyses were participants with prevalent cancer at recruitment ( $\mathrm{n}=27,264$ ); missing data on body size measurements ( $\mathrm{n}=3032$ ); prevalent type-2 diabetes or unknown diabetes status at recruitment (as diabetes medications can affect circulating levels of IGF proteins ${ }^{18}$; $n=26,698$ ); those who reported oral contraceptive and menopausal hormone use at recruitment (as oral estrogens exert a strong first-pass effect on the liver that alters hepatic protein production and changes circulating levels of multiple hormones, including IGF-system markers ${ }^{18}$; $n=19,802$ ); and participants without an IGF1 measurement ( $n=28,480$ ). Our analysis therefore included 397,380 participants.

Blood collection and laboratory methods. As part of the UK Biobank Biomarker Project, ${ }^{19}$ serum levels of IGF1 (DiaSorin Liaison XL), testosterone, and sex hormone binding globulin (SHBG) (all Beckman Coulter DXI 800) were determined by a chemiluminescent immunoassay. The Immunoturbidimetric method (Beckman Coulter DXI 800) was used to assay serum high sensitivity C-reactive protein (CRP) levels. Glycated hemoglobin (HbA1c) levels were determined using the high-performance liquid chromatography Variant II Turbo 2.0 system (Bio-Rad, Hercules, CA). Full details on assay performance have been published. ${ }^{19}$ In summary, average withinlaboratory (total) coefficient of variation for low, medium,
and high internal quality control level samples for each biomarker ranged from $1.7 \%$ to $15.3 \%$ (for IGF1, the coefficients of variation ranged from $5.3 \%$ to $6.2 \%$ ). ${ }^{19}$ A total of 16,357 participants had IGF1 levels measured in blood samples collected at both the recruitment and repeat assessment visit (median of 4 years apart).

Assessment of outcome. Incident cancer cases and cancer cases recorded first in death certificates within the UK Biobank cohort were identified through linkage to national cancer and death registries. Complete follow-up was available through March 31, 2016, for England and Wales and October 31, 2015, for Scotland. Cancer incidence data were coded using the 10th Revision of the International Classification of Diseases (ICD-10). Proximal colon cancers included those found within the cecum, appendix, ascending colon, hepatic flexure, transverse colon, and splenic flexure (C18.0-18.5). Distal colon cancers included those found within the descending (C18.6) and sigmoid (C18.7) colon. Overlapping (C18.8) and unspecified (C18.9) lesions of the colon were included in colon cancers only. Cancer of the rectum included cancers occurring at the recto-sigmoid junction (C19) and rectum (C20).

Statistical analysis. To assess reproducibility between the 2 measurements of IGF1 available in a subsample of participants, we calculated intraclass correlation coefficients (ICC) by dividing the between-person variance by the sum of the between-person and within-person variances.

Hazard ratios (HRs) and 95\% confidence intervals (CIs) were estimated using Cox proportional hazards models. Age was the primary time variable in all models. Time at entry was age at recruitment. Exit time was age at whichever of the following came first: colorectal cancer diagnosis, death, or the last date at which follow-up was considered complete. Models were stratified by age at recruitment in 5 -year categories, Townsend deprivation index quintiles, and region of the recruitment assessment centre. Deviations from proportionality were assessed using an analysis of Schoenfeld residuals, ${ }^{20}$ with no evidence of nonproportionality detected. IGF1 was modeled on the continuous scale (per 1-standard deviation [SD] increment of log-IGF1) and with participants grouped into sexspecific quintiles of circulating levels. Continuous scale models were additionally corrected for regression dilution using regression dilution ratios obtained from participants with repeated IGF1 measurement. ${ }^{21,22}$ To obtain corrected HRs, the $\log$ HRs and their standard errors were divided by the regression dilution ratio for IGF1 (0.76) and then exponentiated. ${ }^{23}$

The multivariable model (model 1) was adjusted for a set of a priori-determined colorectal cancer risk factors, namely waist circumference, total physical activity, height, alcohol consumption frequency, smoking status and intensity, frequency of red and processed meat consumption, family history of colorectal cancer, educational level, regular aspirin/ ibuprofen use, and ever use of hormone replacement therapy. We also additionally adjusted the multivariable models (model 2) for markers of inflammation, sex hormones, and glycemic pathways that are known to interrelate/cross talk with IGFsystem markers, ${ }^{18}$ and have been linked to colorectal cancer risk in some studies, ${ }^{24-27}$ namely CRP, testosterone, SHBG, and HbA1c. Statistical tests for trend were calculated using the ordinal quintile IGF1 entered into the model as a continuous variable. Analyses were also conducted by sex and anatomic
subsite (colon, proximal colon, distal colon, and rectal cancer). Heterogeneity of associations by sex and across subsite was assessed by calculating $\chi^{2}$ statistics. Possible nonlinear effects were modeled using restricted cubic spline models with 5 knots placed at Harrell's default percentiles of circulating IGF1 levels. ${ }^{28}$

The associations between circulating IGF1 and colorectal cancer were further assessed across subgroups of body mass index (BMI; $<25, \geq 25 \mathrm{~kg} / \mathrm{m}^{2}$ ), height (below or above median), age at recruitment ( $<60$ years, $\geq 60$ years), follow-up time ( $<5$ years, $\geq 5$ years), smoking status (never, former, current), and circulating levels (below or above median) of CRP, HbA1c, testosterone, and SHBG. Interaction terms (multiplicative scale) between these variables and circulating IGF1 levels were included in separate models and the statistical significance of the cross-product terms were evaluated using the likelihood ratio test. In a sensitivity analysis, we excluded those participants with circulating levels of HbA1c $\geq 48 \mathrm{mmol} / \mathrm{mol}$ (or $6.5 \%$, the cutoff for type-2 diabetes).

Statistical tests were all 2 -sided and $P<0.05$ was considered statistically significant.

## Mendelian Randomization

We conducted 2-sample MR analyses, in which 2 different independent study samples (GWAS) were used to estimate the single nucleotide polymorphism (SNP)-risk factor (circulating IGF1 and IGFBP3 levels) and SNP-outcome (colorectal cancer) associations to allow causal effect estimates of risk factoroutcome association to be obtained.

Genetic determinants of IGF1 and IGFBP3. Genetic markers for circulating IGF1 and IGFBP3 levels comprised SNPs that were identified ( $P<5 \times 10^{-8}$ ) from the largest GWAS to date. ${ }^{29,30}$ For IGF1, this GWAS was of 358,072 participants from the UK Biobank. ${ }^{29}$ The GWAS analyses of IGFBP3 combined data on 18,995 individuals from 13 studies. ${ }^{30}$ All participants were of European ancestry. From the genome-wide significant variants identified in these GWAS, we excluded correlated SNPs based on a linkage disequilibrium level of $R^{2}<0.01$. Consequently, the instruments for IGF1 ( 413 SNPs spanning 22 autosomes) and IGFBP3 ( 4 SNPs spanning 3 autosomes) explained 9.4\% (F-statistic 89.9) and 6.1\% (F-statistic 308.4) of variability in circulating levels, respectively. Summary information on the genetic instruments and the effect estimates for each individual SNP regarding their association with IGF1 and IGFBP3 levels are presented in Supplementary Tables 1 and 2.

Data on colorectal cancer. Summary data for the associations of the IGF1- and IGFBP3-related genetic variants with colorectal cancer were obtained from a GWAS of 99,152 participants ( 52,865 colorectal cancer cases and 46,287 controls). The GWAS data were from a meta-analysis within the ColoRectal Transdisciplinary Study, the Colon Cancer Family Registry, and the Genetics and Epidemiology of Colorectal Cancer consortium. ${ }^{15}$ Imputation was performed using the Haplotype Reference Consortium r1.0 reference panel and the regression models were further adjusted for age, sex, genotyping platform (whenever appropriate), and genomic principal components as detailed here. ${ }^{31}$ Strength-of-association estimates for each individual SNP with colorectal cancer are presented in Supplementary Table 2.

Statistical power. The a priori statistical power was calculated using an online tool at http://cnsgenomics.com/
shiny $/ \mathrm{mRnd} /{ }^{32}$ Given a type 1 error of $5 \%$, we had sufficient power ( $>80 \%$ ) to detect an odds ratio (OR) per 1 SD for colorectal cancer risk of $\leq 0.94 / \geq 1.06$ for IGF1 and $\leq 0.93$ / $\geq 1.08$ for IGFBP3.

Statistical analysis. A 2 -sample MR approach using summary data and a likelihood-based approach was implemented. Likelihood-based MR analyses are considered the most accurate method to estimate causal effects when there is a continuous log-linear association between risk factor and disease risk. The causal estimate of $X$ on $Y(\beta)$, assumed to be the same for all genetic variants $(k)$, is obtained from the likelihood function of the following model ${ }^{33,34}$ :

$$
\begin{gathered}
X_{k} \sim \mathscr{N}\left(\xi_{k}, \sigma_{X k}^{2}\right) \\
Y_{k} \sim \mathscr{N}\left(\beta_{L} \xi_{k}, \sigma_{Y k}^{2}\right) \text { for } k=1, \ldots, K .
\end{gathered}
$$

MR results correspond to an OR per 1-SD increment in genetically predicted circulating levels of IGF1 and IGFBP3. Heterogeneity of associations by sex and across colorectal anatomic subsites was assessed by calculating $\chi^{2}$ statistics. Cochran's Q statistics quantified heterogeneity across the individual SNPs. Sensitivity analyses were used to check and correct for the presence of pleiotropy in the causal estimates. To evaluate the extent to which directional pleiotropy (nonbalanced horizontal pleiotropy in the MR risk estimates) may have affected the causal estimates for the IGF1 and colorectal cancer association, we used an MR-Egger regression approach. ${ }^{35}$ We also computed OR estimates using the complementary weightedmedian method that can give valid MR estimates under the presence of horizontal pleiotropy when up to $50 \%$ of the included instruments are invalid. ${ }^{36}$ The MR-PRESSO (Mendelian Randomization Pleiotropy RESidual Sum and Outlier) distortion test was used to estimate if horizontal pleiotropy caused by any identified outlier SNPs biased the effect estimates $(P<.05) .{ }^{37}$ As a visual evaluation of pleiotropy, we also provided funnel plots depicting the weight exerted on the effect estimate along the $y$ axis and estimates of the effect on colorectal cancer along the $x$ axis for each SNP used in the corresponding instrumental variable. For the IGFBP3 instrument that had 4 SNPs, we conducted leave-one-out analyses to assess the influence of individual variants on the observed associations.

## Results

## UK Biobank: Serologic Analysis

After a median follow-up time of 7.1 years, 2665 cases of colorectal cancer (1539 in men and 1126 in women) were recorded. Compared with those in the lowest quintile, participants in the highest circulating IGF1 quintile were younger, taller, had lower BMI and waist circumference, were more likely to be never smokers, and less likely to be daily or almost daily consumers of alcohol and to be regular aspirin/ibuprofen users (Table 1). In addition, participants in the highest circulating IGF1 quintile had lower circulating CRP, testosterone, SHBG, and HbA1c levels.

The reproducibility (ICC) of IGF1 levels measured at both the recruitment and repeat assessment visit ( $\mathrm{n}=$ 16,357 participants; median of 4 years apart) was 0.78 (95\% CI 0.77-0.79).

Association between circulating IGF1 levels and colorectal cancer risk. For men and women combined, higher circulating IGF1 levels were associated with elevated colorectal cancer risk in the multivariable models (model 1, HR comparing quintile 5 vs 1 [q5-q1] = 1.24, 95\% CI 1.101.40; $P$-trend $=.006$ ) (Table 2). This positive association strengthened after additional adjustment for circulating levels of CRP, HbA1c, testosterone, and SHBG (model 2, HR $[\mathrm{q} 5-\mathrm{q} 1]=1.34,95 \%$ CI $1.18-1.52 ; P$-trend $<0.0001$ ). In the restricted cubic spline model, no deviation from linearity for the relationship between IGF1 and colorectal cancer was observed ( $P$-nonlinear $=0.91$ ). In the continuous multivariable model 2 , adjusted for circulating levels of CRP, HbA1c, testosterone, and SHBG, a 1-SD increment of IGF1 was associated with an 8\% higher colorectal cancer risk (HR 1.08; 95\% CI 1.04-1.13). Correction for regression dilution resulted in a stronger positive association (HR per 1-SD increment of IGF1, HR 1.11; 95\% CI 1.05-1.17) (Table 2; Figure 1). Similar magnitude positive associations were found for men and women ( $P$-heterogeneity $=.1$ ), and across anatomic subsites (colon vs rectal, $P$-heterogeneity $=$

1; proximal colon vs distal colon, $P$-heterogeneity $=0.8$ ) (Table 2).

The positive association observed between circulating IGF1 levels and colorectal cancer was evident across subgroups of BMI, height, age at recruitment/blood collection, follow-up time, smoking status, and circulating levels of CRP, HbA1c, testosterone, and SHBG (all $P$-interactions $\geq .1$ ) (Figure 1). In a sensitivity analysis, similar relationships for circulating IGF1 levels and colorectal cancer were found when participants with HbA1c levels $\geq 48 \mathrm{mmol} / \mathrm{mol}$ (or $6.5 \%$, cutoff for type-2 diabetes), were excluded from the analyses (Supplementary Table 3).

## Mendelian Randomization

Association between genetically determined circulating IGF1 levels and colorectal cancer risk. We estimated that a $1-S D$ increment in genetically determined IGF1 levels was associated with an $8 \%$ higher colorectal cancer risk (OR 1.08; 95\% CI 1.03-1.12; $P$ value $=3.3 \times 10^{-4}$ ) (Table 3). Positive associations of similar magnitude were

Table 1.Characteristics of UK Biobank Study Participants by Category of Circulating IGF1 levels ( $\mathrm{n}=397,380$ participants

| Baseline characteristic | IGF1 levels |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Q1 | Q2 | Q3 | Q4 | Q5 |
| IGF1, $\mathrm{nmol} / \mathrm{L}^{\text {a }}$ | 14.3 (2.2) | 18.6 (1.1) | 21.4 (0.9) | 24.1 (1.0) | 29.5 (3.9) |
| Colorectal cancer ( n cases) | 595 | 586 | 528 | 475 | 481 |
| Age at recruitment, $y^{\text {a }}$ | 59.1 (7.2) | 57.6 (7.6) | 56.3 (7.9) | 55.1 (8.2) | 53.2 (8.3) |
| Women (\%) | 55.5 | 52.5 | 52.5 | 52.5 | 52.5 |
| Body mass index ( $\left.\mathrm{kg} / \mathrm{m}^{2}\right)^{\text {a }}$ | 28.3 (5.4) | 27.4 (4.7) | 27.1 (4.4) | 26.8 (4.2) | 26.6 (4.0) |
| Waist circumference (cm) ${ }^{\text {a }}$ | 92.7 (14.2) | 90.4 (13.1) | 89.5 (12.7) | 88.9 (12.4) | 88.2 (12.2) |
| Height (cm) ${ }^{\text {a }}$ | 167.7 (9.2) | 168.4 (9.3) | 168.9 (9.3) | 169.2 (9.3) | 169.8 (9.3) |
| Total physical activity (MET hours per wk) (\%) |  |  |  |  |  |
| <10 | 23.8 | 21.8 | 21.1 | 21.2 | 21.0 |
| $\geq 60$ | 22.8 | 23.1 | 22.7 | 22.1 | 21.3 |
| Smoking status (\%) |  |  |  |  |  |
| Never | 50.3 | 53.3 | 55.5 | 57.2 | 59.5 |
| Current | 11.8 | 10.9 | 10.2 | 10 | 9.8 |
| Alcohol consumption (\%) |  |  |  |  |  |
| Never | 9.7 | 7.4 | 6.9 | 6.6 | 6.7 |
| Daily or almost daily | 24.2 | 22.7 | 21.4 | 19.4 | 16.2 |
| Socioeconomic status (Townsend deprivation index) (\%) |  |  |  |  |  |
| Highest quintile | 22.2 | 19.2 | 18 | 18 | 17.8 |
| Family history (first degree relative) of colorectal cancer (\%) |  |  |  |  |  |
| Yes | 11.5 | 11.2 | 10.7 | 10.5 | 10.1 |
| Regular aspirin/ibuprofen use (\%) |  |  |  |  |  |
| Yes | 26.8 | 25.6 | 24.8 | 24.8 | 24.8 |
| Red and processed meat (\%) |  |  |  |  |  |
| $<1$ occasion per week | 10.3 | 10 | 9.9 | 9.9 | 10.2 |
| $\geq 3$ occasions per week | 23.4 | 22.3 | 21.8 | 21.2 | 20.7 |
| Ever menopausal hormone use (\%) ${ }^{b}$ |  |  |  |  |  |
| Yes | 43.8 | 38.1 | 33.8 | 30 | 24.9 |
| CRP ( $m g / L)^{a}$ | 3.4 (5.1) | 2.6 (4.1) | 2.3 (3.8) | 2.1 (3.7) | 1.8 (3.8) |
| Testosterone ( $\mathrm{nmol} / \mathrm{L})^{\text {a }}$ | 7.0 (6.3) | 6.9 (6.2) | 6.8 (6.1) | 6.7 (6.1) | 6.6 (6.0) |
| SHBG ( $\mathrm{nmol} / \mathrm{L})^{\text {a }}$ | 55.4 (28.6) | 52.2 (25.8) | 50.5 (24.9) | 49.1 (24.2) | 46.5 (23.3) |
| $\mathrm{HbA1c}(\mathrm{mmol} / \mathrm{mol})^{2}$ | 35.9 (5.3) | 35.4 (4.5) | 35.1 (4.4) | 34.9 (4.2) | 34.7 (4.5) |

[^1]Table 2.Risk (HRs) of Colorectal Cancer Associated With Circulating IGF1 Levels in the UK Biobank

Table 2.Continued

|  | Q1 | Q2 | Q3 | Q4 | Q5 | $P$-trend | HR per 1-SD increment | HR per 1-SD increment (adjusted) ${ }^{2}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Model ${ }^{2}$ | 1 | 1.26 (1.03-1.54) | 1.24 (1.01-1.53) | 1.23 (0.99-1.53) | 1.33 (1.06-1.66) | 0.03 | 1.08 (1.01-1.16) | 1.11 (1.01-1.22) |
| Men |  |  |  |  |  |  |  |  |
| Model ${ }^{2}$ | 1 | 1.13 (0.89-1.44) | 1.15 (0.90-1.48) | 1.09 (0.84-1.42) | 1.30 (1.00-1.69) | 0.11 | 1.06 (0.97-1.16) | 1.08 (0.96-1.21) |
| Women |  |  |  |  |  |  |  |  |
| Model ${ }^{2}$ | 1 | 1.58 (1.11-2.26) | 1.48 (1.02-2.15) | 1.58 (1.08-2.32) | 1.43 (0.94-2.16) | 0.12 | 1.13 (1.00-1.28) | 1.18 (1.00-1.38) |

NOTE. Model ${ }^{1}$ : Multivariable Cox regression model using age as the underlying time variable and stratified by sex, Townsend deprivation index (quintiles), region of the recruitment assessment center, and age at recruitment. Models adjusted for waist circumference (per 5 cm ), total physical activity ( $<10,10$ to $<20,20$ to $<40,40$ to $<60$, $\geq 60$ metabolic equivalent hours per week, unknown), height (per 10 cm ), alcohol consumption frequency (never, special occasions only, 1-3 times per month, 1-2 times per week, 3-4 times per week, daily/almost daily, unknown), smoking status and intensity (never, former, current- $<15$ per day, current- $\geq 15$ per day, current- intensity unknown, unknown), frequency of red and processed meat consumption ( $<2,2$ to $<3,3$ to $<4, \geq 4$ occasions per week, unknown), family history of colorectal cancer (no, yes, unknown), educational level (Certificates of secondary education [CSEs]/Ordinary [O]-levels/General Certificates of Secondary Education [GCSEs] or equivalent, equivalent, other professional qualifications, college/university degree, none of the above, unknown), regular aspirin/ibuprofen use (no, yes, unknown), ever use of hormone

[^2]found for men and women when analyzed separately $(P$ heterogeneity $=.24$ ), and by anatomic subsite (colon vs rectal, $P$-heterogeneity $=1$; proximal colon versus distal colon, $P$-heterogeneity $=.55$ ) (Table 3). Evidence of effect heterogeneity Cochran's $Q P$ values $<.001$ ) and horizontal pleiotropy (MR-PRESSO $P$ values $<1 \times 10^{-4}$ ) were observed for each model (Supplementary Table 4), but there was little evidence of directional pleiotropy (MREgger intercept $P$ values $>.20$ ) for all models, except for distal colon cancer (MR-Egger intercept $P$ value $=.02$ ). However, the corrected estimates for distal colon cancer and all other models from the MR-Egger and the weightedmedian approach analyses replicated the initial positive effect estimates (Table 3; Supplementary Table 4), and the MR-PRESSO test identified a few outlier SNPs that did not distort the effect estimates (MR-PRESSO $P$ values $>.8$ ). Finally, the funnel plot for the IGF1 instruments indicated a symmetric distribution of effect estimates (Supplementary Figure 1).

Association between genetically determined circulating IGFBP3 levels and colorectal cancer risk. A 1-SD increment in genetically determined IGFBP3 levels was associated with a $12 \%$ higher colorectal cancer risk (OR 1.12; 95\% CI 1.06-1.18; $P=4.2 \times 10^{-5}$ ) (Table 3). Similar positive associations were found for men and women when analyzed separately ( $P$-heterogeneity $=1$ ), and by anatomic subsite (colon vs rectal, $P$-heterogeneity $=$ .87; proximal colon vs distal colon, $P$-heterogeneity $=.77$ ) (Table 3). Evidence of effect heterogeneity was weak ( $P>$ .10) for all models except colon cancer and distal colon cancer models ( $P$ values .04 and .02 , respectively). The MRPRESSO test identified possible horizontal pleiotropy for colon cancer and distal colon cancer ( $P<.05$ ). However, little evidence of directional pleiotropy was found for all models (MR-Egger intercept $P>.15$ ), and the estimates from the weighted-median approach and MR-Egger were consistent with those of likelihood-based approach (Table 3; Supplementary Table 4). The leave-one-out analysis revealed that the positive association between IGFBP3 and colorectal cancer was driven by 1 variant, rs11977526, located in the IGFBP3 gene region (OR per SD increment in IGFBP3 with rs11977526 excluded $=1.02$; 95\% CI $0.92-$ 1.13; $P=.7$ ) (Supplementary Table 5).

## Discussion

In serologic analyses of UK Biobank data, we found that higher prediagnostic levels of circulating IGF1 were associated with greater colorectal cancer risk, with similarstrength positive associations found by sex and colorectal anatomical subsite. Concordant with this, the MR analyses revealed a positive effect of IGF1 on colorectal cancer risk, suggesting that higher circulating IGF1 levels may have a causal role in the development of this malignancy. Collectively, these findings provide strong support for a role of the IGF-pathway in colorectal tumorigenesis.

Despite extensive experimental evidence indicating a role for IGF1 in colorectal tumorigenesis, results from previous epidemiological studies (most including <500
colorectal cancer cases) have been inconsistent, with null or weak positive associations found that did not meet the statistical significance threshold. ${ }^{6-11}$ Our study, using UK Biobank data, is unique in that rather than following a nested study design including a subset of the cohort, IGF1 levels were measured in all study participants. Further, we were able to control statistically for other serologic risk factors that are related to both circulating IGF1 levels and colorectal cancer risk (CRP, testosterone, SHBG, and HbA1c). Our study was the largest to date to examine the IGF1 and
colorectal cancer relationship (including $>2600$ incident cases) which meant we were sufficiently powered to study the IGF1 and colorectal cancer relationship by sex, anatomic subsite of tumor, and across subgroups of other risk factors. We found no heterogeneity for the positive association by sex, anatomic subsite, follow-up time, and other colorectal cancer risk factors.

Despite the robustness of this positive IGF1 and colorectal cancer association, as with all observational studies, these analyses are vulnerable to the inherent biases of this


Table 3.MR Estimates Between Circulating Levels of IGF1 and IGFBP3 and Risk of Colorectal Cancer ( $\mathrm{n}=52,865$ Colorectal Cancer Cases and $n=46,287$ Controls)

|  | MR |  |  |
| :---: | :---: | :---: | :---: |
|  | Likelihood-based approach | Weighted median approach | MR-Egger test |
|  | OR (95\% CI) per 1-SD increment | OR (95\% CI) per 1-SD increment | OR ( $95 \% \mathrm{Cl}$ ) per 1-SD increment |
| IGF1 |  |  |  |
| Colorectal cancer |  |  |  |
| Both sexes | 1.08 (1.03-1.12) | 1.12 (1.04-1.20) | 1.11 (0.98-1.27) |
| Men | 1.11 (1.05-1.17) | 1.07 (0.97-1.18) | 1.18 (1.00-1.38) |
| Women | 1.05 (0.99-1.11) | 1.07 (0.97-1.18) | 1.06 (0.89-1.25) |
| Colon cancer |  |  |  |
| Both sexes | 1.06 (1.01-1.11) | 1.12 (1.03-1.23) | 1.14 (0.98-1.32) |
| Proximal colon cancer |  |  |  |
| Both sexes | 1.04 (0.97-1.11) | 1.08 (0.97-1.21) | 1.02 (0.84-1.23) |
| Distal colon cancer |  |  |  |
| Both sexes | 1.07 (1.00-1.14) | 1.10 (0.98-1.23) | 1.26 (1.05-1.50) |
| Rectal cancer |  |  |  |
| Both sexes | 1.06 (1.00-1.14) | 1.01 (0.90-1.13) | 1.09 (0.91-1.30) |
| IGFBP3 |  |  |  |
| Colorectal cancer |  |  |  |
| Both sexes | 1.12 (1.06-1.18) | 1.14 (1.07-1.20) | 1.18 (1.07-1.31) |
| Men | 1.12 (1.04-1.21) | 1.13 (1.04-1.22) | 1.14 (0.98-1.34) |
| Women | 1.12 (1.04-1.21) | 1.13 (1.04-1.22) | 1.25 (1.18-1.32) |
| Colon cancer |  |  |  |
| Both sexes | 1.11 (1.04-1.19) | 1.13 (1.05-1.21) | 1.27 (1.16-1.38) |
| Proximal colon cancer |  |  |  |
| Both sexes | 1.12 (1.03-1.22) | 1.12 (1.02-1.23) | 1.25 (1.21-1.29) |
| Distal colon cancer |  |  |  |
| Both sexes | 1.10 (1.01-1.20) | 1.13 (1.03-1.24) | 1.29 (1.01-1.65) |
| Rectal cancer |  |  |  |
| Both sexes | 1.12 (1.04-1.22) | 1.14 (1.05-1.25) | 1.18 (1.01-1.39) |

study design, namely residual confounding and reverse causality. MR analyses are less susceptible to such biases because alleles are randomly assigned during meiosis and germline genetic variants are unaffected by disease process. ${ }^{38}$ Our MR analyses yielded strikingly similar estimates of effect for IGF1 and colorectal cancer to our UK Biobank serologic analyses. Similar to the serologic analyses we found no heterogeneity by sex or anatomic subsite in our MR analyses. An important assumption of MR is that the
genetic variants do not influence the outcome through a pathway independent of the main exposure of interest (horizontal pleiotropy). We conducted multiple sensitivity analyses to test for the influence of pleiotropy on our causal estimates, and our results for IGF1 were robust according to these various tests.

Experimental data support a role for IGF1 and its downstream signaling pathways in colorectal tumorigenesis. IGF1 can promote cellular proliferation through the

Figure 1. Subgroup analyses of the association between circulating IGF1 levels and colorectal cancer risk in the UK Biobank. HR per 1-SD increment in circulating IGF1 levels. Multivariable Cox regression model using age as the underlying time variable and stratified by sex, Townsend deprivation index (quintiles), region of the recruitment assessment center, and age at recruitment. Models adjusted for waist circumference (per 5 cm ), total physical activity ( $<10,10$ to $<20,20$ to $<40,40$ to $<60$, $\geq 60$ metabolic equivalent hours per week, unknown), height (per 10 cm ), alcohol consumption frequency (never, special occasions only, 1-3 times per month, 1-2 times per week, 3-4 times per week, daily/almost daily, unknown), smoking status and intensity (never, former, current- $<15$ per day, current- $\geq 15$ per day, current- intensity unknown, unknown), frequency of red and processed meat consumption ( $<2,2$ to $<3,3$ to $<4$, $\geq 4$ occasions per week, unknown), family history of colorectal cancer (no, yes, unknown), educational level (Certificates of secondary education [CSEs]/Ordinary [O]-levels/General Certificates of Secondary Education [GCSEs] or equivalent, National Vocational Qualification [NVQ]/Higher National Diploma [HND]/ Higher National Certificate [HNC]/Advanced [A]-levels/Advanced Subsidiary [AS]-levels or equivalent, other professional qualifications, college/university degree, none of the above, unknown), regular aspirin/ibuprofen use (no, yes, unknown), ever use of hormone replacement therapy (no, yes, unknown), and circulating levels (sex-specific quintiles, missing/unknown) of CRP ( $\mathrm{mg} / \mathrm{L}$ ), HbA1c ( $\mathrm{mmol} / \mathrm{mol}$ ), testosterone ( $\mathrm{nmol} / \mathrm{L}$ ), and SHBG ( $\mathrm{nmol} / \mathrm{L}$ ). ${ }^{\text {a HRs per SD increment were corrected for }}$ regression dilution using a regression dilution ratio ( 0.76 ) obtained from the subsample of participants with repeat IGF1 measurements. Median values: height $=176 \mathrm{~cm}$ for men and 162 cm for women; CRP $=1.3 \mathrm{mg} / \mathrm{L}$; HbA1c $=35 \mathrm{mmol} / \mathrm{mol}$; testosterone $=1 \mathrm{nmol} / \mathrm{L}$ for women and $11.8 \mathrm{nmol} / \mathrm{L}$ for men; $\mathrm{SHBG}=56.3 \mathrm{nmol} / \mathrm{L}$ for men and $37.2 \mathrm{nmol} / \mathrm{L}$ for women.
activation of the mitogen activated protein kinase and the phosphoinositide 3-kinase pathways. ${ }^{39}$ Colonocytes express IGF1-receptors and these are frequently overexpressed in neoplastic cells. ${ }^{40}$ IGF1 has been shown to promote cellular proliferation in colorectal tissue, and the blockade of the IGF1-receptor by a monoclonal antibody inhibits cell proliferation. ${ }^{41}$

Potentially modifiable determinants of circulating IGF1 levels include dietary protein intake, ${ }^{42-44} \mathrm{BMI},{ }^{43}$ and physical activity. ${ }^{45}$ In addition, a number of pharmacologic agents targeting the IGF system have been developed, although, as of yet, they have not been successful in treating colorectal cancer patients in clinical trials. ${ }^{46}$ Although circulating IGF1 levels are modifiable, it is currently unknown how long an intervention aimed at altering IGF1 concentrations would have to be applied to have measurable effects. Our result provides possible evidence for a causal role of circulating IGF1 levels in colorectal cancer development and will hopefully reinvigorate efforts to develop and introduce interventions targeting the IGF system for colorectal cancer prevention in susceptible individuals.

The bioactivity of IGF1 is partially regulated through IGFBPs, with most bound to IGFBP3. Higher levels of IGFBP3 both increase the serologic binding capacity for IGF1, but also lower circulating IGF1 bioavailability. ${ }^{47}$ As well as IGF-binding properties, IGFBP3 has been shown to induce apoptosis and reduce proliferation in colon cancer cell lines. ${ }^{4,5}$ The positive association we found in our MR analyses for IGFBP3 and colorectal cancer is inconsistent with its anticipated anti-tumorigenic effects. Epidemiologic studies examining the IGFBP3 and colorectal cancer association have reported mixed findings, with inverse, positive, and null results all previously reported. ${ }^{6,7,10-14}$ The positive effect estimate in our MR analysis was driven solely by 1 variant, rs11977526, with a null result found when this variant was excluded. Although rs11977526 is located in close proximity to the IGFBP3 gene, this variant has also been associated with several anthropometric traits ${ }^{48}$ and circulating levels of IGF2, ${ }^{49}$ for which there is evidence of a role in colorectal cancer development. The IGF2 gene is imprinted and loss of imprinting has been detected in colorectal cancer tumor tissue. ${ }^{50}$ Stromainduced IGF2 has been shown to promote colon cancer progression in a paracrine and autocrine manner. ${ }^{51} \mathrm{~A}$ meta-analysis of 3 prospective studies reported an almost 2 -fold higher colorectal cancer risk when the highest and lowest groups of circulating IGF2 levels were compared. ${ }^{52}$ MR analyses for IGF2 are currently limited as a GWAS for circulating IGF2 has not been conducted. Similarly, for IGFBP3 and other IGF system components (including bioavailable IGF1 levels), more precise genetic instruments are now required to help disentangle the possible biological effects of specific ligands and binding proteins on colorectal cancer development.

This was the largest and most comprehensive study that used 2 complementary study designs to examine the role of IGF1 on colorectal cancer development. A limitation of our analysis is that IGF1 levels were measured only
once at baseline in all participants, and it is possible that these measurements may not reflect exposure levels across time. However, our reproducibility analysis in a subset of the cohort found an ICC value of 0.78 between IGF1 measurements collected 4 years apart, indicating that a single measurement provides a good estimate of medium to longer-term exposures. The availability of repeat IGF1 measurements also allowed us to correct for regression dilution bias, and thus limiting the effects of measurement error and within-person variability on our risk estimates ${ }^{21}$; the positive associations we found were stronger after regression dilution correction. A limitation of our MR analyses is that our use of summary-level data precluded the investigation of nonlinear effects and subgroup analyses by other risk factors (eg, age, BMI, smoking). However, our serologic analysis, which yielded a similar overall risk estimate to our MR result, found no evidence of a nonlinear association for IGF1 and colorectal cancer and little heterogeneity across subgroups of other risk factors.

In conclusion, our complementary serologic and MR analyses provide strong support for a positive association of circulating IGF1 levels on colorectal cancer risk. This result suggests that diet/lifestyle ${ }^{42-44,53}$ or pharmacologic ${ }^{1}$ interventions targeting the IGF system may offer a promising strategy in reducing the risk of colorectal cancer.

## Supplementary Material

Note: To access the supplementary material accompanying this article, visit the online version of Gastroenterology at www.gastrojournal.org, and at https://doi.org/ 10.1053/j.gastro.2019.12.020.

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A


OR per SD increment in insulin-like growth factor-1 (IGF-1) levels

## B



OR per SD increment in insulin-like growth factor-binding protein-3 (IGFBP-3) levels

Supplementary Table 1.Summary Information on the IGF1 and IGFBP3 Genetic Instruments Used in the Mendelian Randomization Analyses

| Exposure | Source/Publication | n SNPs | Variance explained (\%) | Mean (SD) |
| :--- | :--- | :---: | :---: | :---: |
| IGF1 | UK Biobank/Sinnott-Armstrong et al 2019 | 413 | 9.4 | $21.6(5.6) \mathrm{nmol} / \mathrm{L}$ |
| IGFBP3 | Meta-analysis/Teumer et al $2016^{2}$ | 4 | 6.1 | $3340.4(863.3) \mathrm{ng} / \mathrm{mL}$ |

Supplementary Figure 1. Funnel plots of risk estimates of $(A)$ IGF1 and $(B)$ IGFBP3 with colorectal cancer against instrumental strength. Instrumental strength is SNP to colorectal cancer effect corrected by SNP to IGF1 or IGFBP3 standard error of the effect. $X$-axis is in logarithmic scale. $P$ values are 2 -sided, MR test.

Supplementary Table 2.Association Parameters of Instrumental SNPs Used in the IGF1 and IGFBP3 Genetic Instruments

| SNP | Chr | Position | Gene ${ }^{\text {a }}$ | Effect Other allele allele |  | Association parameters for the exposure |  | Association parameters for colorectal cancer |  | Association parameters for colorectal cancer men |  | Association parameters for colorectal cancer women |  | Association parameters for colon cancer |  | Association parameters for proximal colon cancer |  | Association parameters for distal colon cancer |  | Association parameters for rectal cancer |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE |
| IGF1-n |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| SNPs $=413$ |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| rs903908 | , | 2,202,967 | SKI | T | C | -0.016 | 0.003 | -0.006 | 0.009 | -0.014 | 0.013 | 0.003 | 0.013 | -0.016 | 0.011 | -0.006 | 0.014 | -0.023 | 0.014 | 0.016 | 0.014 |
| rs17393144 | 1 | 9,210,262 | MIR34A | G | A | -0.016 | 0.003 | 0.006 | 0.010 | 0.019 | 0.013 | -0.007 | 0.014 | 0.005 | 0.011 | 0.012 | 0.015 | -0.002 | 0.015 | -0.005 | 0.015 |
| rs112436634 | 1 | 10,637,709 | PEX14 | C | T | -0.016 | 0.003 | 0.001 | 0.009 | -0.014 | 0.013 | 0.014 | 0.013 | -0.002 | 0.011 | -0.014 | 0.015 | 0.018 | 0.015 | 0.017 | 0.015 |
| rs17037452 | 1 | 11,895,675 | CLCN6 | A | G | 0.023 | 0.003 | 0.000 | 0.012 | -0.011 | 0.017 | 0.011 | 0.018 | 0.000 | 0.015 | 0.029 | 0.019 | -0.029 | 0.019 | -0.003 | 0.019 |
| rs36086195 | 1 | 16,510,894 | ARHGEF19 -AS1 | T | C | -0.019 | 0.003 | 0.018 | 0.009 | 0.012 | 0.013 | 0.028 | 0.013 | 0.020 | 0.011 | 0.003 | 0.014 | 0.043 | 0.014 | 0.010 | 0.014 |
| rs12723255 | 1 | 21,233,570 | EIF4G3 | T | C | -0.017 | 0.003 | 0.002 | 0.009 | 0.004 | 0.013 | 0.001 | 0.013 | 0.003 | 0.011 | -0.002 | 0.014 | 0.008 | 0.014 | 0.010 | 0.014 |
| rs6701954 | 1 | 22,022,176 | USP48 | T | G | 0.014 | 0.003 | 0.005 | 0.009 | 0.007 | 0.012 | -0.001 | 0.013 | -0.008 | 0.011 | 0.007 | 0.014 | -0.022 | 0.014 | 0.014 | 0.014 |
| rs76914895 | 1 | 23,292,603 | LACTBL1 | T | C | -0.027 | 0.005 | 0.046 | 0.019 | 0.054 | 0.026 | 0.038 | 0.027 | 0.030 | 0.022 | 0.026 | 0.029 | 0.054 | 0.030 | 0.093 | 0.029 |
| rs2075995 | 1 | 23,847,464 | E2F2 | A | C | -0.014 | 0.003 | 0.003 | 0.009 | 0.011 | 0.012 | -0.003 | 0.013 | 0.008 | 0.011 | -0.004 | 0.014 | 0.024 | 0.014 | 0.014 | 0.014 |
| rs2802330 | 1 | 26,466,831 | PDIK1L | A | G | -0.031 | 0.003 | 0.003 | 0.012 | -0.001 | 0.016 | 0.004 | 0.016 | 0.011 | 0.014 | 0.025 | 0.018 | -0.009 | 0.018 | -0.004 | 0.018 |
| rs17360994 | 1 | 27,278,573 | KDF1 | C | T | -0.042 | 0.005 | -0.012 | 0.017 | -0.003 | 0.023 | -0.023 | 0.024 | -0.006 | 0.020 | -0.012 | 0.026 | 0.022 | 0.026 | -0.005 | 0.026 |
| rs569356 | 1 | 29,136,686 | OPRD1 | A | G | -0.027 | 0.004 | 0.007 | 0.013 | -0.016 | 0.018 | 0.035 | 0.019 | 0.005 | 0.016 | 0.019 | 0.020 | -0.001 | 0.021 | 0.012 | 0.020 |
| rs3131646 | 1 | 40,383,552 | MYCL | G | T | 0.016 | 0.003 | -0.006 | 0.010 | -0.006 | 0.013 | -0.008 | 0.014 | -0.013 | 0.012 | -0.029 | 0.015 | 0.002 | 0.015 | 0.002 | 0.015 |
| rs61780439 | 1 | 41,490,177 | SLFNL1-AS1 | G | A | 0.021 | 0.003 | 0.004 | 0.011 | 0.012 | 0.015 | -0.001 | 0.015 | -0.011 | 0.013 | -0.013 | 0.017 | -0.007 | 0.017 | 0.013 | 0.017 |
| rs2819336 | 1 | 44,015,809 | PTPRF | C | T | -0.027 | 0.003 | -0.004 | 0.009 | -0.003 | 0.013 | -0.006 | 0.013 | -0.007 | 0.011 | -0.019 | 0.014 | 0.001 | 0.015 | 0.008 | 0.014 |
| rs7539178 | 1 | 65,383,002 | JAK1 | A | C | -0.026 | 0.004 | -0.024 | 0.013 | -0.026 | 0.018 | -0.023 | 0.018 | -0.036 | 0.016 | -0.047 | 0.020 | -0.008 | 0.021 | -0.015 | 0.020 |
| rs1046011 | 1 | 65,898,996 | LEPR / LEPROT | C | T | -0.021 | 0.003 | 0.011 | 0.010 | 0.006 | 0.014 | 0.017 | 0.014 | 0.005 | 0.012 | -0.001 | 0.015 | 0.014 | 0.015 | 0.015 | 0.015 |
| rs1430753 | 1 | 68,692,642 | WLS | G | A | -0.021 | 0.003 | -0.011 | 0.011 | -0.012 | 0.016 | -0.010 | 0.016 | -0.007 | 0.013 | 0.000 | 0.017 | -0.020 | 0.018 | -0.041 | 0.017 |
| rs165316 | 1 | 91,533,297 | RPL5P6 | A | G | -0.073 | 0.003 | $-0.018$ | 0.011 | -0.045 | 0.015 | 0.009 | 0.016 | -0.037 | 0.013 | -0.035 | 0.017 | -0.033 | 0.017 | 0.006 | 0.017 |
| rs1825813 | 1 | 92,708,973 | C1orf146 | G | A | -0.023 | 0.003 | -0.027 | 0.011 | -0.044 | 0.016 | -0.009 | 0.016 | -0.037 | 0.013 | -0.034 | 0.017 | -0.040 | 0.018 | -0.019 | 0.017 |
| rs599839 | 1 | 109,822,166 | PSRC1/ CELSR2 | A | G | -0.031 | 0.003 | 0.006 | 0.011 | 0.013 | 0.015 | -0.002 | 0.015 | 0.009 | 0.013 | -0.004 | 0.016 | 0.022 | 0.017 | 0.029 | 0.016 |
| rs45505697 | 1 | 153,651,058 | NPR1 | C | A | 0.030 | 0.005 | 0.011 | 0.020 | -0.009 | 0.028 | 0.037 | 0.029 | 0.021 | 0.024 | 0.015 | 0.031 | 0.023 | 0.032 | 0.022 | 0.031 |
| rs1127313 | 1 | 154,556,425 | ADAR | G | A | 0.024 | 0.003 | 0.009 | 0.009 | 0.007 | 0.012 | 0.013 | 0.013 | 0.007 | 0.011 | -0.001 | 0.014 | 0.017 | 0.014 | 0.019 | 0.014 |
| rs77369503 | 1 | 163,027,266 | RGS4 | G | A | 0.045 | 0.007 | -0.004 | 0.028 | -0.002 | 0.038 | -0.013 | 0.041 | -0.027 | 0.033 | -0.048 | 0.042 | -0.051 | 0.043 | 0.012 | 0.042 |
| rs75681856 | 1 | 174,916,323 | RABGAP1L | C | T | -0.023 | 0.004 | 0.006 | 0.013 | 0.016 | 0.019 | -0.007 | 0.019 | -0.004 | 0.016 | -0.010 | 0.021 | -0.007 | 0.021 | 0.015 | 0.021 |
| rs12749024 | 1 | 176,522,365 | PAPPA2 | C | T | -0.075 | 0.004 | -0.008 | 0.013 | -0.019 | 0.018 | 0.000 | 0.018 | -0.011 | 0.015 | -0.008 | 0.020 | -0.009 | 0.020 | -0.013 | 0.020 |
| rs10913351 | 1 | 177,447,742 | AL122019.1 | G | A | -0.032 | 0.005 | 0.021 | 0.018 | 0.028 | 0.026 | 0.017 | 0.026 | 0.020 | 0.022 | 0.022 | 0.028 | 0.023 | 0.029 | 0.004 | 0.028 |
| rs11577063 | 1 | 179,341,999 | AXDND1 | G | T | -0.020 | 0.003 | -0.011 | 0.011 | 0.007 | 0.015 | -0.031 | 0.015 | -0.011 | 0.013 | -0.025 | 0.016 | -0.002 | 0.017 | -0.020 | 0.016 |
| rs143885630 | 1 | 183,482,785 | SMG7 | G | A | 0.030 | 0.004 | -0.013 | 0.014 | -0.013 | 0.019 | -0.008 | 0.020 | -0.014 | 0.017 | -0.045 | 0.022 | 0.016 | 0.022 | 0.019 | 0.022 |
| rs940400 | 1 | 200,269,134 | LINC00862 | C | A | 0.025 | 0.004 | 0.012 | 0.014 | 0.016 | 0.019 | 0.007 | 0.020 | 0.023 | 0.017 | 0.001 | 0.021 | 0.014 | 0.022 | -0.016 | 0.021 |
| rs7545345 | 1 | 205,690,941 | NUCKS1 | T | C | -0.026 | 0.004 | 0.000 | 0.013 | 0.007 | 0.019 | -0.007 | 0.019 | 0.002 | 0.016 | -0.017 | 0.021 | 0.016 | 0.021 | -0.013 | 0.021 |
| rs2724373 | 1 | 207,999,200 | C1orf132 | C | T | 0.019 | 0.003 | -0.014 | 0.009 | -0.003 | 0.013 | -0.030 | 0.013 | -0.024 | 0.011 | -0.024 | 0.014 | -0.022 | 0.015 | -0.011 | 0.014 |
| rs10779509 | 1 | 209,728,370 | AL023754.1 | T | C | -0.014 | 0.003 | -0.012 | 0.009 | -0.024 | 0.013 | 0.000 | 0.013 | -0.016 | 0.011 | -0.013 | 0.014 | -0.021 | 0.014 | -0.017 | 0.014 |
| rs340837 | 1 | 214,162,734 | PROX1 | T | G | -0.021 | 0.003 | $-0.025$ | 0.009 | -0.034 | 0.012 | -0.011 | 0.013 | -0.032 | 0.011 | -0.029 | 0.014 | -0.029 | 0.014 | -0.019 | 0.014 |
| rs12141189 | 1 | 221,053,545 | HLX | C | T | -0.045 | 0.003 | 0.023 | 0.010 | 0.015 | 0.015 | 0.033 | 0.015 | 0.031 | 0.013 | 0.033 | 0.016 | 0.027 | 0.017 | 0.000 | 0.016 |
| rs4306136 | 1 | 221,608,720 | AL360013.2 | A | G | 0.017 | 0.003 | 0.014 | 0.009 | 0.001 | 0.013 | 0.029 | 0.013 | 0.010 | 0.011 | 0.026 | 0.014 | 0.001 | 0.014 | 0.022 | 0.014 |
| rs708108 | 1 | 228,189,855 | WNT3A | C | T | -0.015 | 0.003 | -0.007 | 0.009 | -0.009 | 0.013 | -0.009 | 0.013 | -0.007 | 0.011 | -0.013 | 0.014 | -0.001 | 0.014 | -0.011 | 0.014 |
| rs684818 | 1 | 234,854,779 | AL160408.6 | T | C | 0.024 | 0.003 | 0.030 | 0.009 | 0.028 | 0.012 | 0.029 | 0.013 | 0.021 | 0.011 | 0.019 | 0.014 | 0.027 | 0.014 | 0.038 | 0.014 |



| SNP | Chr | Position | Gene ${ }^{\text {a }}$ | Effect Other allele allele |  | Association parameters for the exposure |  | Association parameters for colorectal cancer |  | Association parameters for colorectal cancer men |  | Association parameters for colorectal cancer women |  | Association parameters for colon cancer |  | Association parameters for proximal colon cancer |  | Association parameters for distal colon cancer |  | Association parameters for rectal cancer |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE | Effect | SE |
| rs7267595 | 20 | 10,643,850 | JAG1 | A | C | 0.015 | 0.003 | 0.005 | 0.009 | 0.023 | 0.012 | -0.012 | 0.013 | -0.004 | 0.011 | -0.015 | 0.014 | 0.006 | 0.014 | 0.012 | 0.014 |
| rs2273058 | 20 | 20,033,319 | CRNKL1 | G | A | -0.022 | 0.003 | -0.005 | 0.009 | -0.002 | 0.012 | -0.008 | 0.013 | -0.011 | 0.011 | -0.013 | 0.014 | -0.009 | 0.014 | -0.002 | 0.014 |
| rs6106324 | 20 | 20,964,988 | AL133465.1 | T | C | 0.019 | 0.003 | -0.011 | 0.009 | -0.009 | 0.013 | -0.012 | 0.013 | -0.008 | 0.011 | -0.011 | 0.014 | -0.007 | 0.014 | -0.010 | 0.014 |
| rs2424396 | 20 | 21,630,280 | LINC01726 | A | G | -0.033 | 0.004 | 0.016 | 0.016 | 0.003 | 0.022 | 0.026 | 0.022 | 0.025 | 0.019 | 0.034 | 0.024 | 0.022 | 0.025 | 0.020 | 0.024 |
| rs6088579 | 20 | 33,284,624 | PIGU/NCOA6 | G | A | 0.027 | 0.003 | -0.016 | 0.012 | -0.021 | 0.017 | -0.011 | 0.017 | -0.010 | 0.014 | -0.016 | 0.019 | -0.004 | 0.019 | -0.045 | 0.019 |
| rs2207132 | 20 | 39,142,516 | MAFB | G | A | 0.048 | 0.007 | 0.059 | 0.028 | 0.033 | 0.039 | 0.090 | 0.042 | 0.033 | 0.034 | 0.046 | 0.045 | 0.000 | 0.045 | 0.059 | 0.045 |
| rs17265513 | 20 | 39,832,628 | ZHX3 | C | T | -0.022 | 0.003 | -0.016 | 0.011 | -0.029 | 0.016 | -0.001 | 0.016 | -0.018 | 0.014 | -0.014 | 0.017 | -0.017 | 0.018 | -0.005 | 0.017 |
| rs16995311 | 20 | 49,201,102 | PTPN1 | A | C | 0.040 | 0.005 | 0.024 | 0.018 | 0.036 | 0.024 | 0.010 | 0.025 | 0.015 | 0.021 | 0.006 | 0.027 | 0.027 | 0.028 | 0.032 | 0.027 |
| rs2104476 | 20 | 54,852,856 | - | A | G | -0.020 | 0.003 | 0.006 | 0.010 | 0.012 | 0.014 | -0.001 | 0.014 | 0.003 | 0.012 | 0.006 | 0.015 | 0.003 | 0.016 | -0.004 | 0.015 |
| rs2738787 | 20 | 62,328,375 | RTEL <br> 1- <br> TNFR <br> SF6B/ <br> RTEL1 | G | A | -0.037 | 0.005 | -0.028 | 0.017 | -0.060 | 0.023 | 0.007 | 0.024 | -0.005 | 0.020 | 0.001 | 0.026 | 0.000 | 0.026 | -0.064 | 0.025 |
| rs9978775 | 21 | 40,694,526 | BRWD1-AS1 | G | A | 0.019 | 0.003 | 0.009 | 0.009 | 0.021 | 0.012 | -0.001 | 0.013 | 0.008 | 0.011 | 0.008 | 0.014 | 0.001 | 0.014 | 0.011 | 0.014 |
| rs8138950 | 22 | 29,448,643 | ZNRF3 | C | T | 0.015 | 0.003 | 0.016 | 0.009 | 0.013 | 0.012 | 0.020 | 0.013 | 0.005 | 0.011 | -0.004 | 0.014 | 0.010 | 0.014 | 0.032 | 0.014 |
| rs2412973 | 22 | 30,529,631 | HORMAD2 | C | A | -0.014 | 0.003 | 0.024 | 0.009 | 0.011 | 0.012 | 0.035 | 0.013 | 0.028 | 0.011 | 0.037 | 0.014 | 0.022 | 0.014 | 0.016 | 0.014 |
| rs12106594 | 22 | 31,885,316 | DRG1/ <br> EIF4 <br> ENIF1/ <br> SFI1 | C | T | -0.036 | 0.006 | 0.007 | 0.020 | -0.017 | 0.028 | 0.030 | 0.029 | 0.008 | 0.024 | 0.010 | 0.031 | -0.004 | 0.031 | 0.003 | 0.031 |
| rs5755948 | 22 | 36,179,095 | RBFOX2 | G | A | -0.028 | 0.004 | 0.003 | 0.013 | 0.008 | 0.018 | -0.005 | 0.019 | 0.011 | 0.016 | 0.028 | 0.020 | -0.005 | 0.021 | -0.001 | 0.020 |
| rs6519133 | 22 | 39,096,602 | JOSD1 | T | C | 0.029 | 0.003 | 0.011 | 0.009 | 0.012 | 0.013 | 0.011 | 0.013 | 0.005 | 0.011 | 0.013 | 0.014 | 0.002 | 0.015 | 0.024 | 0.014 |
| rs9611565 | 22 | 41,767,486 | TEF | T | C | 0.029 | 0.003 | 0.001 | 0.010 | 0.010 | 0.014 | -0.010 | 0.015 | 0.006 | 0.012 | -0.001 | 0.016 | 0.006 | 0.016 | -0.015 | 0.016 |
| rs4823324 | 22 | 46,238,123 | ATXN10 | T | C | 0.016 | 0.003 | 0.002 | 0.009 | 0.001 | 0.013 | 0.003 | 0.013 | 0.006 | 0.011 | 0.023 | 0.014 | -0.013 | 0.014 | 0.000 | 0.014 |
| IGFBP3 n SNPs = 4 |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| rs4234798 | 4 | 7,219,933 | SORCS2 | G | T | 0.095 | 0.011 | 0.012 | 0.009 | 0.020 | 0.013 | 0.003 | 0.013 | 0.005 | 0.011 | -0.003 | 0.014 | 0.015 | 0.014 | 0.015 | 0.014 |
| rs11977526 | 7 | 46,008,110 | IGFBP3 | A | G | 0.287 | 0.011 | 0.037 | 0.009 | 0.034 | 0.013 | 0.042 | 0.013 | 0.042 | 0.011 | 0.042 | 0.014 | 0.042 | 0.014 | 0.037 | 0.014 |
| rs700753 | 7 | 46,753,684 | TNS3 | G | C | 0.158 | 0.011 | -0.006 | 0.009 | -0.006 | 0.013 | -0.007 | 0.013 | -0.009 | 0.011 | 0.001 | 0.014 | -0.025 | 0.015 | -0.010 | 0.014 |
| rs1065656 | 16 | 1,838,836 | NUBP2 | G | C | 0.111 | 0.011 | 0.010 | 0.010 | 0.015 | 0.013 | 0.006 | 0.014 | -0.003 | 0.012 | 0.004 | 0.015 | -0.010 | 0.015 | 0.017 | 0.015 |

[^3]Supplementary Table 3.Risk (HRs) of Colorectal Cancer Associated With Circulating IGF1 Levels With Those Participants With Circulating HbA1c Levels $\geq 48 \mathrm{mmol} / \mathrm{mol}$ (or $6.5 \%$, the Cutoff for Type 2 Diabetes) Excluded

| Q1 | Q2 | Q3 | Q4 | Q5 | $P$-trend | HR per 1-SD increment | HR per 1-SD increment (adjusted) $^{a}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Colorectal cancer |  |  |  |  |  |  |  |
| Both sexes 1 | 1.13 (1.00-1.27) | 1.14 (1.01-1.29) | 1.12 (0.98-1.27) | 1.34 (1.18-1.53) | $<0.0001$ | 1.09 (1.04-1.14) | 1.12 (1.06-1.18) |
| Colon cancer |  |  |  |  |  |  |  |
| Both sexes 1 | 1.06 (0.91-1.22) | 1.09 (0.94-1.26) | 1.06 (0.91-1.24) | 1.34 (1.15-1.57) | 0.002 | 1.08 (1.03-1.14) | 1.11 (1.04-1.19) |
| Proximal colon cancer |  |  |  |  |  |  |  |
| Both sexes 1 | 1.15 (0.94-1.40) | 1.22 (1.00-1.50) | 0.97 (0.77-1.21) | 1.35 (1.08-1.68) | 0.09 | 1.08 (1.00-1.16) | 1.10 (1.00-1.21) |
| Distal colon cancer |  |  |  |  |  |  |  |
| Both sexes 1 | 0.96 (0.77-1.20) | 0.97 (0.77-1.23) | 1.18 (0.94-1.48) | 1.31 (1.03-1.65) | 0.01 | 1.09 (1.00-1.18) | 1.12 (1.01-1.24) |
| Rectal cancer |  |  |  |  |  |  |  |
| Both sexes 1 | 1.29 (1.05-1.59) | 1.27 (1.02-1.57) | 1.25 (1.00-1.56) | 1.36 (1.08-1.72) | 0.02 | 1.10 (1.02-1.19) | 1.14 (1.03-1.25) |

NOTE. Multivariable Cox regression model using age as the underlying time variable and stratified by sex, Townsend deprivation index (quintiles), region of the recruitment assessment center, and age at recruitment. Models adjusted for waist circumference (per 5 cm ), total physical activity ( $<10,10$ to $<20,20$ to $<40,40$ to $<60, \geq 60$ metabolic equivalent hours per week, unknown), height (per 10 cm ), alcohol consumption frequency (never, special occasions only, 1-3 times per month, 1-2 times per week, 3-4 times per week, daily/almost daily, unknown), smoking status and intensity (never, former, current- <15 per day, current- $\geq 15$ per day, current- intensity unknown, unknown), frequency of red and processed meat consumption ( $<2$, 2 to $<3,3$ to $<4, \geq 4$ occasions per week, unknown), family history of colorectal cancer (no, yes, unknown), educational level (Certificates of secondary education [CSEs]/Ordinary [O]-levels/General Certificates of Secondary Education [GCSEs] or equivalent, National Vocational Qualification [NVQ]/Higher National Diploma [HND]/Higher National Certificate [HNC]/ Advanced [A]-levels/Advanced Subsidiary [AS]-levels or equivalent, other professional qualifications, college/university degree, none of the above, unknown), regular aspirin/ibuprofen use (no, yes, unknown), ever use of hormone replacement therapy (no, yes, unknown), and circulating levels (sex-specific quintiles, missing/unknown) of CRP (mg/L), HbA1c (mmol/mol), testosterone ( $\mathrm{nmol} / \mathrm{L}$ ), and SHBG ( $\mathrm{nmol} / \mathrm{L}$ ).
${ }^{2}$ HRs per SD increment were additionally corrected for regression dilution using a regression dilution ratio (0.76) obtained from the subsample of participants with repeat IGF1 measurements.

Supplementary Table 4.MR Pleiotropy Sensitivity Tests for the Circulating IGF1 and IGFBP3 Levels and Risk of Colorectal Cancer


LCI , lower confidence interval; UCI, upper confidence interval.

Supplementary Table 5.MR leave-one-out analysis for IGFBP3 Levels and Risk of Colorectal Cancer

|  | MR |  |
| :--- | :---: | :---: |
|  | Likelihood-based approach | $P$ value |
| SNP excluded | OR (95\% CI) per 1-SD increment | 0.7 |
| rs11977526 | $1.02(0.92-1.13)$ | $1 . E-03$ |
| rs4234798 | $1.11(1.04-1.18)$ | $2 . E-05$ |
| rs900750 | $1.15(1.08-1.23)$ | $4 . E-04$ |


[^0]:    ${ }^{1}$ Section of Nutrition and Metabolism, International Agency for Research on Cancer, Lyon, France; ${ }^{2}$ Colorectal Cancer Group, ONCOBELL Program, Bellvitge Biomedical Research Institute (IDIBELL), L'Hospitalet de Llobregat, Barcelona, Spain; ${ }^{3}$ Division of Gastroenterology, Massachusetts General Hospital and Harvard Medical School, Boston, Massachusetts; ${ }^{4}$ Clinical and Translational Epidemiology Unit, Massachusetts General Hospital and Harvard Medical School, Boston, Massachusetts; ${ }^{5}$ Department of Epidemiology, Harvard T.H. Chan School of Public Health, Harvard University, Boston, Massachusetts;
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[^1]:    MET, metabolic equivalents.
    ${ }^{2}$ Mean and SD.
    ${ }^{b}$ Among women only.

[^2]:    replacement therapy (no, yes, unknown).
    Model $^{2}$ : Model
    Model ${ }^{2}$ : Model ${ }^{1}$ plus additional adjustment for circulating levels (sex-specific quintiles, missing/unknown) of CRP ( $\mathrm{mg} / \mathrm{L}$ ), $\mathrm{HbA1c}(\mathrm{mmol} / \mathrm{mol})$, testosterone ( $\mathrm{nmol} / \mathrm{L}$ ), and
    SHBG (nmol/L).
    ${ }^{\text {a }}$ HRs per SD increment were additionally corrected for regression dilution using a regression dilution ratio ( 0.76 ) obtained from the subsample of participants with repeat IGF1 measurements.

[^3]:    Chr, chromosome; SE, standard error.
    ${ }^{\text {a }}$ Overlapped or nearest gene (sourced from: https://snp-nexus.org/index.html). Where blank, gene unknown.

