

Dimensionality Reduction and Uncertainty Quantification for Inverse Problems

Tristan van Leeuwen - Utrecht University, the Netherlands

SIAM CSE - MS35 - March 14, 2015

Inverse problems

Estimate parameters from noisy measurements

$$\mathbf{d}_i = F(\mathbf{m})\mathbf{q}_i + \mathbf{n}_i,$$

with

\mathbf{d}_i - observations

F - forward operator, typically involves a PDE solve

\mathbf{m} - parameters

\mathbf{q}_i - input/source

$\mathbf{n}_i \sim \mathcal{N}(0, C)$

Non-linear data-fitting

Maximum likelihood estimation can be formulated as

$$\min_{\mathbf{m}} \sum_{i=1}^n \|F(\mathbf{m})\mathbf{q}_i - \mathbf{d}_i\|_{C^{-1}}^2,$$

which can be solved using a Newton-like method.

- ▶ Noise covariance may not be known a-priori
- ▶ Evaluation of the misfit and gradient requires $2n$ PDE solves.

Non-linear data-fitting

Agenda:

Estimation of the noise covariance matrix

- ▶ May give different estimate of the parameters \mathbf{m} ,
- ▶ Important for uncertainty quantification

Non-linear data-fitting

Agenda:

Estimation of the noise covariance matrix

- ▶ May give different estimate of the parameters \mathbf{m} ,
- ▶ Important for uncertainty quantification

Reduction of the effective number of simulations

- ▶ Random subsampling
- ▶ Exploit structure of Covariance matrix

Extended least-squares

Formulate an extended LS problem:

$$\min_{\mathbf{m}, C} \log(|C|) + \sum_{i=1}^n \|F(\mathbf{m})\mathbf{q}_i - \mathbf{d}_i\|_{C^{-1}}^2.$$

For fixed \mathbf{m} we have a closed-form solution

$$C(\mathbf{m}) = \sum_{i=1}^n \mathbf{r}_i(\mathbf{m})\mathbf{r}_i(\mathbf{m})^T,$$

where

$$\mathbf{r}_i(\mathbf{m}) = F(\mathbf{m})\mathbf{q}_i - \mathbf{d}_i.$$

[Aravkin et. al. 12]

Intermezzo: Variational projection

Given a twice differentiable function $g(x, y)$, define $\bar{y}(x) = \min_y g(x, y)$ and define a *reduced* function

$$f(x) = g(x, \bar{y}(x)),$$

then

$$\begin{aligned}\nabla f(x) &= \nabla_x g(x, \bar{y}(x)), \\ \nabla^2 f(x) &= \nabla_x^2 g(x, \bar{y}(x)) - \nabla_{x,y}^2 g^T \left(\nabla_y^2 g \right)^{-1} \nabla_{x,y}^2 g.\end{aligned}$$

[Bell & Burke 08; Aravkin et. al. 12]

Extended least-squares

Define a *reduced* objective

$$f(\mathbf{m}) = \log(|C(\mathbf{m})|) + \sum_{i=1}^n \|F(\mathbf{m})\mathbf{q}_i - \mathbf{d}_i\|_{C(\mathbf{m})}^2,$$

with gradient

$$\nabla f(\mathbf{m}) = \sum_{i=1}^n DF(\mathbf{m}, \mathbf{q}_i)^T C(\mathbf{m})^{-1} \mathbf{r}_i(\mathbf{m}).$$

where

$$C(\mathbf{m}) = \sum_{i=1}^n \mathbf{r}_i(\mathbf{m})\mathbf{r}_i(\mathbf{m})^T.$$

Extended least-squares

Main tasks:

1. Compute all residuals (n PDE-solves),

$$\mathbf{r}_i = F(\mathbf{m})\mathbf{q}_i - \mathbf{d}_i,$$

2. Estimate the covariance matrix,

$$C = \sum_{i=1}^n \mathbf{r}_i \mathbf{r}_i^T,$$

3. Compute gradient ($2n$ PDE solves)

$$\nabla f(\mathbf{m}) = \sum_{i=1}^n DF(\mathbf{q}_i)^T C^{-1} \mathbf{r}_i$$

Estimating the covariance

Organizing the residuals in an $m \times n$ matrix R , we have

$$C = RR^T.$$

Expanding $R = U_k \Sigma_k V_k^T$, we have $C = U_k \Sigma_k^2 U_k^T$, and

$$f = n + 2 \sum_{i=1}^k \log(\sigma_i),$$

$$\nabla f = \sum_{i=1}^k \sigma_i^{-1} DF(\tilde{\mathbf{q}}_i)^T \mathbf{v}_i,$$

with $\tilde{Q} = QV_k$.

Estimating the covariance

Observation:

If the covariance matrix has rank k , we need only $2k$ PDE-solves to evaluate the gradient

Question:

How do we efficiently obtain the (truncated) SVD of R ?

Intermezzo: Randomized trace estimation

Given a matrix A , we can estimate the trace

$$\text{tr}(A^T A) \approx \sum_{i=1}^k \mathbf{w}_i^T A^T A \mathbf{w}_i,$$

where \mathbf{w}_i is an i.i.d. random Gaussian vector.

Intermezzo: Randomized trace estimation

Given a matrix A , we can estimate the trace

$$\text{tr}(A^T A) \approx \sum_{i=1}^k \mathbf{w}_i^T A^T A \mathbf{w}_i,$$

where \mathbf{w}_i is an i.i.d. random Gaussian vector.

Such techniques have been very successful in PDE-constrained optimization/inverse problems.

[Avron & Toledo 11; Haber et. al. '12]

Randomized linear algebra

First approach:

1. Compress the matrix by random projection:

$$\tilde{R} = RW_k,$$

where $\mathbb{E}\{W_k W_k^T\} = I_n$.

2. Compute k -SVD: $\tilde{R} = \tilde{U}_k \tilde{\Sigma}_k \tilde{V}_k^T$, and find

$$RR^T \approx \tilde{U}_k \tilde{\Sigma}_k^2 \tilde{U}_k^T.$$

Cost: k PDE solves + k -SVD of $m \times k$ matrix.

Randomized linear algebra

Second approach:

1. Capture range of the matrix $Y = RW_k$ and find orthonormal basis L for Y .
2. Compute k -SVD of $L^T R = \tilde{U}_k \tilde{\Sigma}_k V_k^T$ and find

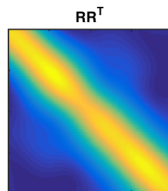
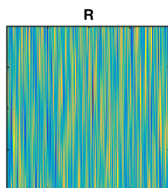
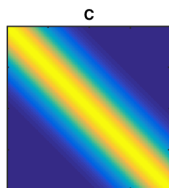
$$RR^T \approx L \tilde{U}_k \tilde{\Sigma}_k^2 V_k^T.$$

Cost: $2k$ PDE-solves + QR of $m \times k$ matrix + k -SVD of $k \times n$ matrix.

[Halko et. al. 11]

Randomized linear algebra

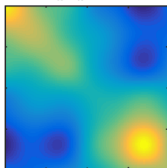
True covariance matrix



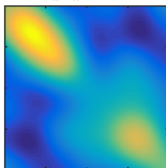
Randomized linear algebra

Stochastic approximation

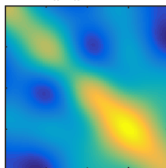
$RW_k W_k^T R^T$ (k=5)



$RW_k W_k^T R^T$ (k=10)



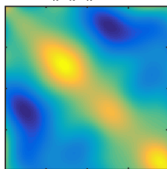
$RW_k W_k^T R^T$ (k=20)



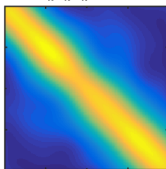
Randomized linear algebra

Randomized SVD

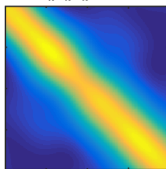
$U_k S_k^2 U_k^T$ (k=5)



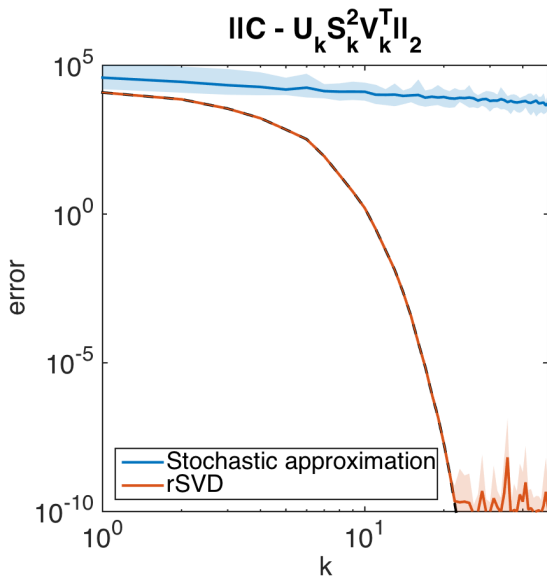
$U_k S_k^2 U_k^T$ (k=10)



$U_k S_k^2 U_k^T$ (k=20)



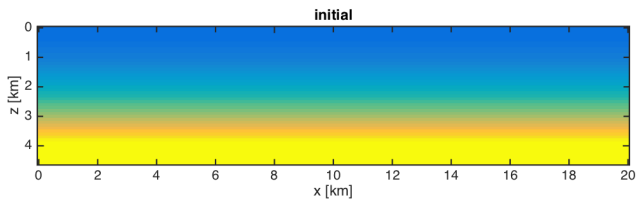
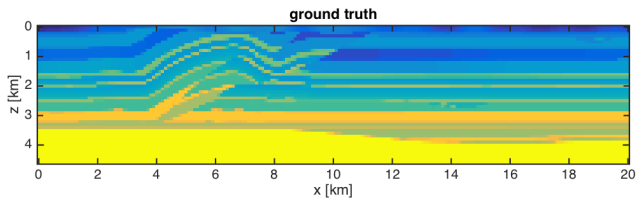
Randomized linear algebra



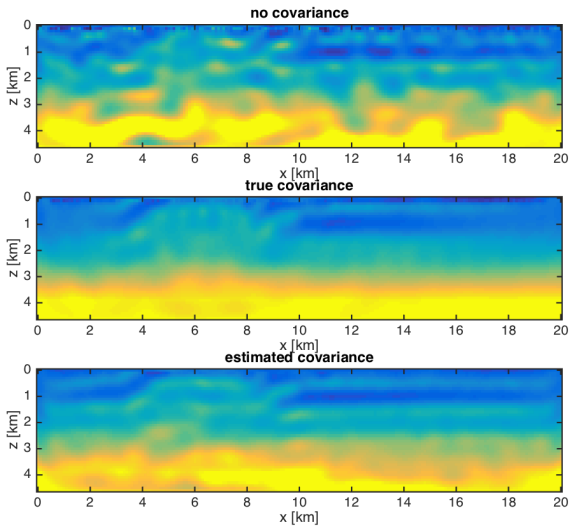
Numerical results

- ▶ PDE: 2D Helmholtz equation
- ▶ gradient-descent with Borzilai-Borwein steplength
- ▶ estimate covariance on-the-fly

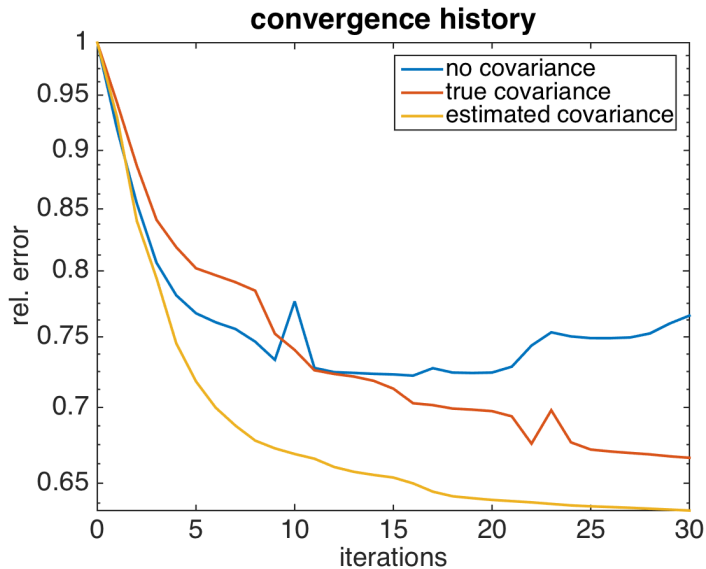
Numerical results



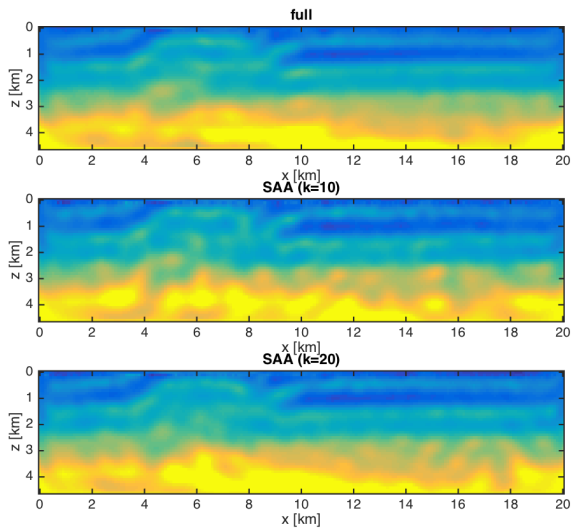
Numerical results



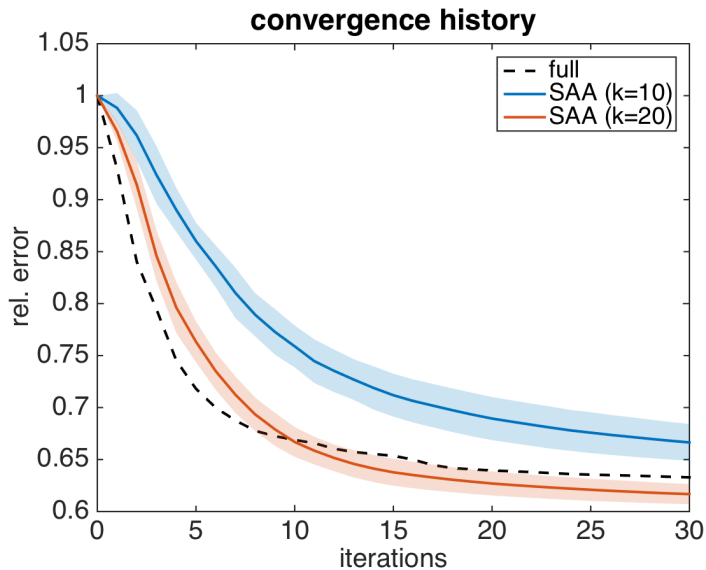
Numerical results



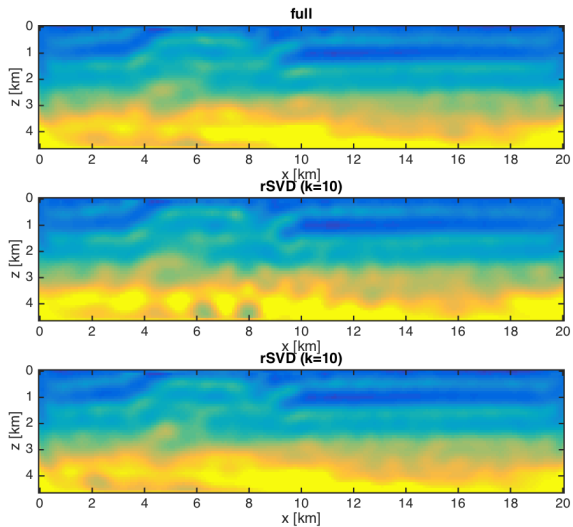
Numerical results



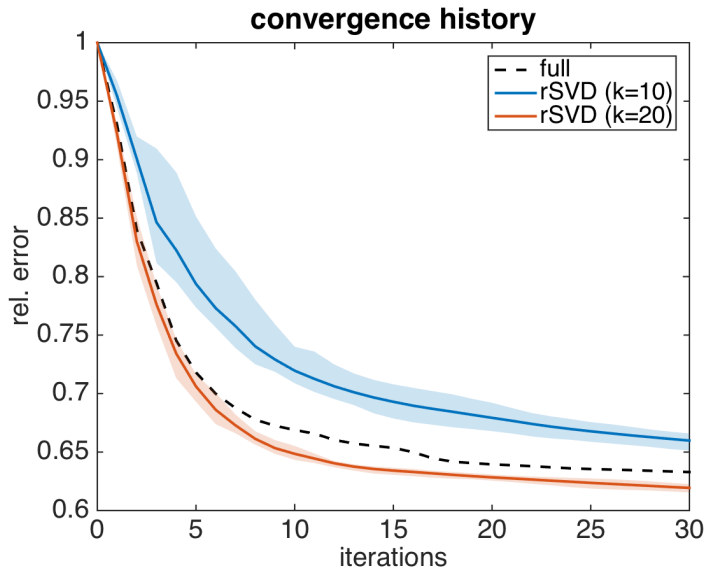
Numerical results



Numerical results



Numerical results



Conclusions

- ▶ We can exploit low-rank structure of the covariance matrix to reduce the # of PDE solves
- ▶ Low rank estimate of the covariance matrix can be computed on-the-fly using stochastic approximation or randomized SVD
- ▶ First results are promising, SA does remarkably well

Future work

- ▶ Adaptive estimation of the rank
- ▶ More sophisticated model for covariance matrix (diagonal + low rank)
- ▶ Exploit low rank structure of C to represent the GN Hessian