

## ARTICLE

# Analysis of Heritability and Shared Heritability Based on Genome-Wide Association Studies for 13 Cancer Types

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## Abstract

**Background:** Studies of related individuals have consistently demonstrated notable familial aggregation of cancer. We aim to estimate the heritability and genetic correlation attributable to the additive effects of common single-nucleotide polymorphisms (SNPs) for cancer at 13 anatomical sites.

**Methods:** Between 2007 and 2014, the US National Cancer Institute has generated data from genome-wide association studies (GWAS) for 49 492 cancer case patients and 34 131 control patients. We apply novel mixed model methodology (GCTA) to this GWAS data to estimate the heritability of individual cancers, as well as the proportion of heritability attributable to cigarette smoking in smoking-related cancers, and the genetic correlation between pairs of cancers.

**Results:** GWAS heritability was statistically significant at nearly all sites, with the estimates of array-based heritability,  $h_1^2$ , on the liability threshold (LT) scale ranging from 0.05 to 0.38. Estimating the combined heritability of multiple smoking characteristics, we calculate that at least 24% (95% confidence interval [CI] = 14% to 37%) and 7% (95% CI = 4% to 11%) of the heritability for lung and bladder cancer, respectively, can be attributed to genetic determinants of smoking. Most pairs of cancers studied did not show evidence of strong genetic correlation. We found only four pairs of cancers with marginally statistically significant correlations, specifically kidney and testes ( $\rho = 0.73$ , SE = 0.28), diffuse large B-cell lymphoma (DLBCL) and pediatric osteosarcoma ( $\rho = 0.53$ , SE = 0.21), DLBCL and chronic lymphocytic leukemia (CLL) ( $\rho = 0.51$ , SE = 0.18), and bladder and lung ( $\rho = 0.35$ , SE = 0.14). Correlation analysis also indicates that the genetic architecture of lung cancer differs between a smoking population of European ancestry and a nonsmoking Asian population, allowing for the possibility that the genetic etiology for the same disease can vary by population and environmental exposures.

**Conclusion:** Our results provide important insights into the genetic architecture of cancers and suggest new avenues for investigation.

Studies of related individuals have consistently demonstrated that there is notable familial aggregation of cancer. The three largest studies, based on the Swedish Family-Cancer Database (1–3), the Utah Population and Cancer Registry Database (4,5), and the Icelandic Cancer Registry (6), have shown familial aggregation for cancer at nearly every anatomical site. For common cancers such as prostate, breast, and lung, the familial relative risk (FRR), defined as the increase in risk associated with each affected first-degree relative of an individual, is generally estimated to be below or around 2.0. In contrast, for some rare cancers occurring early in life, such as those of testes and bone, estimates of FRR can exceed 5. Although shared environmental factors contribute to this aggregation, studies of twins (7,8) and extended family members (6,9,10) have clearly identified a substantial genetic contribution, commonly known as heritability.

Genome-wide association studies (GWAS) have provided an opportunity to study the contribution of common single-nucleotide polymorphisms (SNPs) to the heritability of complex traits, including cancers. In addition to identifying specific susceptibility SNPs, novel mixed-effect modeling methods (11–13) can utilize GWAS data to quantify the additive heritability attributable to all common susceptibility SNPs captured by genotyping arrays, regardless of whether those SNPs individually have reached the stringent level of genome-wide statistical significance. Understanding the total contribution of common SNPs

will be instrumental for evaluating the potential clinical applications of genetics in risk stratification and for guiding future genetic studies of cancer (14,15).

Evidence also suggests that there is overlap between cancers with respect to their genetic architectures. Previous family studies have observed familial co-aggregation of cancers among certain sites (5,6), including pairs of cancers at neighboring sites, such as the colon and rectum, as well as for more distant and seemingly unrelated sites such as the cervix and esophagus (6). GWAS have also directly identified shared regions such as 8q24.1 and 5p15.33 (*TERT-CLPTM1L*) containing SNPs affecting cancer at multiple sites, while earlier studies have identified major genes such as *TP53* (16) and *BRCA* (17) containing highly penetrant rare variants affecting multiple cancers. Sites with overlapping genetic architectures may be studied together to understand shared biology and to increase power to detect susceptibility loci.

In this study, we performed an analysis of heritability and shared heritability for cancer at 13 different sites using data from case/control GWAS of more than 80 000 individuals carried out or reported to the US National Cancer Institute. We expand upon recent GWAS estimates of heritability (13) by nearly doubling the number of cases evaluated, exploring six new cancer (sub)types, and considering populations of non-European ancestry. We use detailed information on multiple

smoking characteristics to assess the proportion of heritability in smoking-related cancers that can be attributed to the genetic determinants of cigarette smoking. Furthermore, we use genetic correlation analysis to assess shared heritability across cancer sites and for lung cancer across two distinct ethnic populations to assess the evidence of gene-environment interactions for this complex malignancy. These analyses provide insights into the contribution of common SNPs to cancer heritability, coheritability and their relationships to a major environmental risk-factor, smoking.

## Methods

### Study Populations

The study includes 49 492 cancer case patients and 34 131 control patients who participated in large case/control GWAS that were genotyped at or reported to the National Cancer Institute's (NCI's) Cancer Genomics Research Laboratory (CGR) between 2007 and 2014. GWAS were conducted in individuals of European ancestry unless otherwise stated. Details about the GWAS can be found in the [Supplementary Methods](#) (available online). Studies were approved by the local institutional review boards, and all participants provided written informed consent.

### Estimation of Heritability and Genetic Correlation

We estimated heritability,  $h_i^2$ , and genetic correlation using mixed model methodology (GCTA) software (11,12), adjusting for sex, substudy, and the top 20 eigenvectors, and after following the stringent quality control procedures described in the [Supplementary Methods](#) (available online). Genetic correlation was considered statistically significant if the two-sided test statistic provided a  $P$  value below .01. To estimate heritability attributable to undiscovered loci, we identified SNPs ([Supplementary Table 1](#), available online) from the National Human Genome Research Institute (NHGRI) catalog or from recent publications (18–20) that were associated with a given cancer ( $P < 5 \times 10^{-8}$ ) and removed all SNPs within 250 KB of those loci prior to calculation of the genetic relation matrix. A description of the GCTA methodology, determining FRR, and sensitivity analyses are provided in the [Supplementary Methods](#) (available online).

### Shared SNP Analysis

Among the 375 SNPs associated with at least one of the studied cancers in the NHGRI catalog or a recent publication, we considered only the 318 SNPs that were outside regions surrounding *TERT* (1280000 +/- 500kb on chromosome 5) and *8q24* (128000000 +/- 500kb on chromosome 8), both already known to have pleiotropic effects on multiple cancers, and that were either genotyped or could be imputed accurately in at least one GWAS. Among those SNPs, we evaluated their associations with all other cancers and reported their associations if the corresponding two-sided  $P$  value was below .001.

Next, we created a polygenic risk score for each cancer. Let “ $T$ ” denote a specific cancer site,  $R_{iT}$  be an individual's risk of cancer “ $T$ ”,  $\Omega_T$  be the subset of the 318 SNPs associated with cancer “ $T$ ”, and  $G_{ij}$  be the number of minor alleles for individual  $i$  at SNP  $j$  in  $\Omega_T$ . Then the polygenic risk score,  $Z_{iT}$ , for individual  $i$  was defined as  $Z_{iT} = \sum_{j \in \Omega_T} \hat{\beta}_{Tj} G_{ij}$ , where  $\hat{\beta}_{Tj}$  was estimated from fitting the model

$$\text{logit}(R_{iT}) \propto \sum_{j \in \Omega_T} \beta_{Tj} G_{ij}$$

using our GWAS. We then tested for an association between the polygenic risk score of the primary cancer and the risk of each of the other cancers and considered the test to be statistically significant if the two-sided  $P$  value was below .05.

### Estimation of the Heritability of Smoking and Its Contribution to Cancer Heritability

To estimate the proportion of cancer heritability that can be explained by genetic determinants of smoking, we first defined a smoking-related risk score variable ( $S_T$ ) that weights three smoking-related risk factors, namely smoking status, intensity, and duration according to their strength of association with a particular cancer ([Supplementary Methods](#), available online). Variability of  $S_T$ , a mathematical construct which we denote by  $\text{Var}(S_T)$ , explains the variation of cancer risk because of smoking in a population on the log-risk scale. The component of  $\text{Var}(S_T)$  attributable to genetics can be estimated as  $V_{GS} = h_{TS}^2 \text{Var}(S_T)$ , where  $h_{TS}^2$  is the heritability of  $S_T$  and  $\text{Var}(S_T)$  is the total variability of  $S_T$  in the underlying population, which we approximate by the 10 530 control patients with relevant smoking variables. The proportion,  $P$ , of cancer heritability attributable to smoking is then estimated as  $P = V_{GS} / V_G$ , where  $V_G = 2 \log(\text{FRR})$  transforms our estimate of cancer heritability to the variability in risk on the log-risk scale (21).

## Results

The estimates of array-based heritability,  $h_i^2$ , on the liability threshold (LT) scale ranged from 0.05 to 0.38 across the 13 cancer sites ([Table 1](#)), with esophageal cancer ( $h_i^2 = 0.38$ , 95% confidence interval [CI] = 0.17 to 0.59, Asian population), prostate cancer ( $h_i^2 = 0.38$ , 95% CI = 0.24 to 0.51), and testicular cancer ( $h_i^2 = 0.30$ , 95% CI = 0.08 to 0.51) displaying the strongest heritable components. After removing known susceptibility loci ([Supplementary Table 1](#), available online), the adjusted estimates of heritability remained similar ([Table 1](#)) for most of the cancer sites, indicating that the majority of the common underlying susceptibility loci remain to be discovered. Known loci constituted the largest proportion of heritability for cancers of the testes (33%) and prostate (25%).

We transformed our measures of the genetic contribution to cancer from heritability on the liability threshold scale to familial relative risk (FRR) for direct comparison with registry-based family studies ([Table 2](#)). For most cancers, the FRR explained by GWAS SNPs was between 1.28 and 1.63. The estimates of FRR were highest for testicular cancer (FRR = 3.09, 95% CI = 1.41 to 6.05), osteosarcoma (2.90, 95% CI = 1.73 to 4.70) and CLL (2.28, 95% CI = 1.86 to 2.77). The GWAS estimates of excess risk (FRR-1) in the studies of European ancestry were in the range of 15% to 53% of the average estimates based on the Icelandic, Swedish, and Utah registry data, with the exception of DLBCL (4.5%). For lung cancer in an Asian population, our observed FRR of 1.31 (95% CI = 1.16 to 1.46) ([Table 2](#)) can be compared with an estimate of 2.44 (95% CI = 1.79 to 3.32) from a Taiwanese family study (22).

Smoking is heritable and is a strong risk factor for both lung and bladder cancers. We estimated the heritability of the lung cancer smoking-related risk score to be 0.15 (95% CI = 0.10 to 0.20) and the heritability of the bladder cancer smoking-related risk score to be 0.16 (95% CI = 0.09 to 0.22) ([Table 3](#); [Supplementary Tables 3–4](#) and [Supplementary Figures 1–2](#), available online).

**Table 1.** GWAS estimates of cancer heritability on the liability scale\*

Cancer	$h_1^2$ (95% CI)	$h_2^2$ (95% CI)
Bladder	0.123 (0.086 to 0.160)	0.112 (0.075 to 0.148)
Breast (ER-)	0.096 (0 to 0.199)	0.079 (0 to 0.181)
Endometrium	0.178 (0.085 to 0.270)	0.177 (0.085 to 0.270)
Esophagus†	0.381 (0.174 to 0.588)	0.370 (0.164 to 0.577)
Glioma	0.046 (0 to 0.116)	0.036 (0 to 0.106)
Kidney	0.147 (0.023 to 0.270)	0.136 (0.012 to 0.259)
Lung		
Asian†,‡	0.121 (0.064 to 0.177)	0.102 (0.046 to 0.159)
European	0.206 (0.142 to 0.271)	0.189 (0.125 to 0.253)
Lymphoma		
CLL	0.220 (0.162 to 0.278)	0.164 (0.105 to 0.222)
DLBCL	0.092 (0.038 to 0.145)	0.088 (0.035 to 0.142)
Osteosarcoma	0.159 (0.079 to 0.239)	0.158 (0.078 to 0.238)
Pancreas	0.098 (0.037 to 0.160)	0.084 (0.023 to 0.145)
Prostate		
Overall	0.378 (0.244 to 0.513)	0.285 (0.151 to 0.419)
Nonadvanced stage	0.351 (0.211 to 0.491)	0.259 (0.120 to 0.399)
Advanced stage	0.232 (0.157 to 0.307)	0.193 (0.119 to 0.268)
Stomach† (noncardia)	0.253 (0 to 0.522)	0.243 (0 to 0.512)
Testes	0.299 (0.084 to 0.513)	0.199 (0 to 0.415)

\*  $h_1^2$  (95% confidence interval [CI]) is the estimated heritability on the liability scale (95% CI) using all qualifying single-nucleotide polymorphisms (SNPs), while  $h_2^2$  is the heritability after removing SNPs within 250 kb of a previous genome-wide association study hit. CI = confidence interval; CLL = chronic lymphocytic leukemia; DLBCL = diffuse large B-cell lymphoma; ER = estrogen receptor; GWAS = genome-wide association study.

† Asian population.

‡ Nonsmoking women.

Based on a comparison with our estimates for the total heritability of cancer (Table 3), we estimated that 23.6% (95% CI = 14.2% to 37.4%) and 7.1% (95% CI = 4.3% to 11.3%) of the heritability for lung and bladder cancer, respectively, can be attributed to the genetic determinants of the three smoking characteristics. When restricting the analysis to ever-smokers, the heritability of the smoking-related risk scores remained similar (Table 3) and we estimated that 22.0% (95% CI = 7.8% to 43.0%) and 3.0% (95% CI = 1.2% to 5.9%) of the total heritability, respectively, was attributable to that of smoking behavior for lung and bladder cancers. Although limited by the number of never-smokers, we also noted lower estimates of heritability in this subgroup (Supplementary Table 5, available online).

The estimated genetic correlations between most pairs of cancer sites were modest and statistically nonsignificant (Figure 1; Supplementary Table 6, available online). Among the 91 compared pairs, all combinations of the 12 solid tumors, CLL, and DLBCL, the median (mean) correlation was 0.031 (0.055) and, in tests for statistical significance, only 12% and 22% of these 91 *P* values were below .1 and .2, providing little evidence of the enrichment that would be expected to occur with many strongly correlated pairs. The four pairs with the strongest correlations ( $P < .01$ ) were kidney and testes ( $\rho = 0.73$ , SE = 0.28), DLBCL and osteosarcoma ( $\rho = 0.53$ , SE = 0.21), DLBCL and CLL ( $\rho = 0.51$ , SE = 0.18), and bladder and lung ( $\rho = 0.35$ , SE = 0.14). The two types, CLL and osteosarcoma, that were individually correlated with DLBCL showed minimal correlation ( $\rho = 0.19$ , SE = 0.15) with each other. For prostate cancer, the best-fitting estimate of the genetic correlation between aggressive and nonaggressive disease was at the upper boundary ( $\rho = 1$ , SE = 0.25), indicating a

**Table 2.** Estimates of first-degree familial relative risk from familial registries and GWAS\*

Cancer	Sweden			Iceland†	Utah†	GWAS
	All 1 <sup>st</sup> -degree relationships FRR (95% CI)	Parent/child FRR (95% CI)	Sibling FRR (95% CI)	FRR (90% CI)	FRR (95% CI)	FRR (95% CI)
Bladder	1.69 (1.33 to 2.14)	1.53 (1.16 to 1.99)	3.30 (1.70 to 5.78)	1.68 (1.39 to 2.05)	1.8 (1.4 to 2.3)	1.37 (1.25 to 1.50)
Breast (ER-)	–	–	–	–	–	1.28 (0.98 to 1.63)
Endometrium	3.02 (2.33 to 3.92)	2.85 (2.08 to 3.82)	3.97 (1.97 to 7.13)	1.86 (1.31 to 2.62)	1.4 (1.1 to 1.8)	1.56 (1.25 to 1.92)
Esophagus	–	2.14 (0.77 to 4.70)	–	2.09 (1.30 to 3.31)	1.3 (0.2 to 10.0)	1.63‡ (1.27 to 2.05)
Glioma	1.67 (1.43 to 1.94)	–	3.31 (2.08 to 5.02)	1.41 (0.74 to 2.40)	2.3 (0.99 to 4.5)	1.19 (0.91 to 1.54)
Kidney	1.78 (1.33 to 2.39)	1.52 (1.06 to 2.11)	4.52 (2.15 to 8.35)	2.30 (1.89 to 2.80)	2.1 (1.3 to 3.5)	1.54 (1.07 to 2.13)
Lung						
European	1.70 (1.42 to 2.05)	1.64 (1.34 to 2.00)	2.61 (1.29 to 4.68)	2.00 (1.83 to 2.16)	2.4 (1.9 to 3.0)	1.42 (1.28 to 1.57)
Asian	–	–	–	–	–	1.31‡,§ (1.16 to 1.46)
Lymphoma						
CLL	8.5 (6.1 to 11.7)	–	–	–	6.1 (4.75 to 7.65)	2.28 (1.86 to 2.77)
DLBCL	9.8 (3.1 to 31.0)	–	–	–	–	1.40 (1.15 to 1.68)
Osteosarcoma	–	–	–	–	–	2.90 (1.73 to 4.70)
Pancreas	–	1.68 (1.16 to 2.35)	–	2.33 (1.83 to 2.96)	2.1 (1.3 to 3.2)	1.35 (1.12 to 1.62)
Prostate	2.75 (2.32 to 3.25)	2.71 (2.26 to 3.22)	4.91 (1.28 to 12.7)	1.89 (1.75 to 2.01)	2.1 (1.9 to 2.2)	1.51 (1.32 to 1.72)
Stomach	1.99 (1.47 to 2.71)	1.72 (1.19 to 2.40)	8.82 (3.50 to 18.3)	1.90 (1.74 to 2.05)	2.0 (1.1 to 3.7)	1.94‡ (0.95 to 3.49)
Testes	7.07 (5.34 to 9.37)	4.31 (2.05 to 7.95)	8.50 (6.01 to 11.7)	3.52 (1.18 to 7.37)	1.8 (0.4 to 8.6)	3.09 (1.41 to 6.05)

\* Comparison of first-degree familial relative risk (95% CI) measured by our genome-wide association study (last column) with estimates of familial relative risk from the three largest family studies. Prior estimates are from (1, 6, 10), except for CLL (31), DLBCL (32), glioma (sibling) (33), pancreas (34), and esophagus (3). CI = confidence interval; CLL = chronic lymphocytic leukemia; DLBCL = diffuse large B-cell lymphoma; ER = estrogen receptor; FRR = family relative risk; GWAS = genome-wide association study.

† All first-degree relationships.

‡ Asian population.

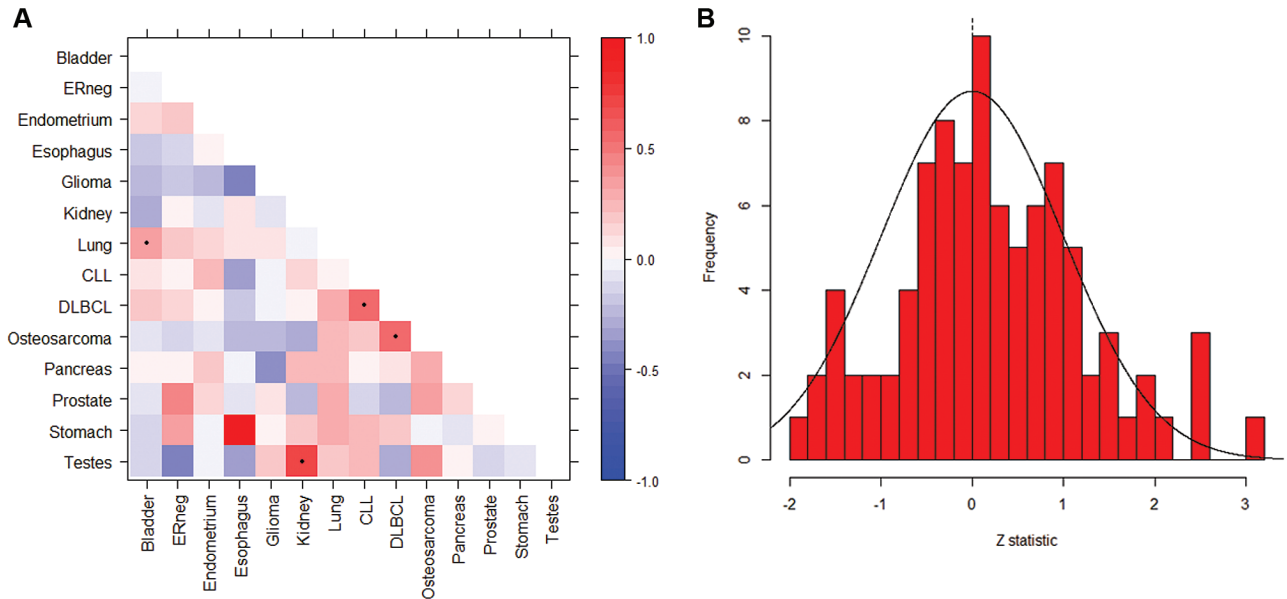
§ Nonsmoking women.



**Table 3.** Estimates of the contribution of smoking to cancer heritability\*

Cancer	All subjects			Former/current smokers		
	$V_{GS}$ : smoking (95% CI)	$V_G$ : cancer (95% CI)	Ratio, % (95% CI)	$V_{GS}$ : smoking (95% CI)	$V_G$ : cancer (95% CI)	Ratio, % (95% CI)
Bladder	0.045 (0.029 to 0.061)	0.62 (0.44 to 0.81)	7.1 (4.3 to 11.3)	0.020 (0.0073 to 0.032)	0.66 (0.41 to 0.88)	3.0 (1.2% to 5.9)
Lung	0.166 (0.098 to 0.23)	0.70 (0.50 to 0.90)	23.6 (14.2 to 37.4)	0.11 (0.043 to 0.167)	0.47 (0.27 to 0.67)	22.0 (7.8% to 43.0)

\*The ratio,  $V_{GS}$ :smoking/ $V_G$ :cancer, estimates the proportion of the genetic variability of cancer that can be attributed to the genetic determinants of smoking, where  $V_{GS}$ :smoking and  $V_G$ :cancer are the genetic variances of the smoking-related risk score and cancer, respectively. CI = confidence interval.



**Figure 1.** Genetic correlation of cancer pairs. **A)** The genetic correlation between cancer sites. **Dots** indicate  $P < .01$ . **B)** Distribution of the corresponding Z-statistics from testing the null hypotheses of no genetic correlations. **Black curve** illustrates the expected distribution under the null hypothesis of no genetic correlation. CLL = chronic lymphocytic leukemia; DLBCL = diffuse large B-cell lymphoma.

shared genetic architecture between the two clinical subtypes of this malignancy. For lung cancer, on the other hand, we estimated the genetic correlation across two distinct racial/ethnic groups and found the correlation between nonsmoking Asian women and individuals of European ancestry to be only modest and statistically nonsignificant ( $\rho = 0.10$ ,  $SE = 0.15$ ), suggesting distinct genetic etiologies for the same malignancy in two distinct populations with different ethnic background and exposure history.

As a second means to assess genetic overlap, we tested whether the 318 SNPs previously associated with one of the 13 cancers at GWAS levels of statistical significance ( $P < 5 \times 10^{-8}$ ), but outside the known pleiotropic regions of 5p15.33 and 8q24, were associated with other cancers (Supplementary Tables 7–8, available online). Among the 318 SNPs, 25 were associated with a second cancer at  $P$  values of less than .001 (Supplementary Tables 7–8, available online). We further calculated a polygenic risk score for each cancer (Methods) with at least ten associated SNPs among the 318 previously discovered. We found individuals with a high risk score for lung cancer were at an increased risk of bladder cancer (RR of bladder cancer comparing 90<sup>th</sup> vs 10<sup>th</sup> percentile = 1.07, 95% CI = 1.05 to 1.10) and individuals with a high risk score for CLL were at an increased risk of DLBCL (RR = 1.12, 95% CI = 1.07 to 1.16) (Supplementary Table 9–12, available online). After adjusting for status, intensity, and duration of smoking, the strength of association between the lung

cancer score and bladder cancer was reduced but remained statistically significant (RR = 1.05, 95% CI = 1.02 to 1.08).

We estimated heritability separately for men and women (Supplementary Tables 13–14, available online). Although we observed differences by sex, none reached statistical significance. When studies were divided by sex, the genetic correlation between cancer in men and women exceeded 0.9 in nearly all scenarios, with the 95% confidence interval always including 1. Moreover, we also estimated heritability under multiple additional analytical settings (Supplementary Table 15, available online) and found the results from these sensitivity analyses to be qualitatively similar. Sensitivity analyses further suggested that population stratification was adequately handled by adjusting for principal components (Supplementary Table 16, available online).

## Discussion

Our analysis confirms that common SNPs meaningfully contribute to risk across the spectrum of cancers and account for different fractions of estimated heritability. The estimates of array-based heritability on the liability threshold scale spanned a wide range, from 5% (glioma) to 38% (prostate, esophageal cancer), and the majority of the observed heritability for most cancers is likely attributable to genetic variants that are not located in previously identified susceptibility regions. The corresponding estimates of familial relative risk were between 1.19 (glioma)

and 2.90 (osteosarcoma). The estimated FRRs of the most common adult cancers were restricted to a narrow range of 1.4 to 1.6, which are lower than those previously observed in family-based studies (1,5,6). A number of additional factors are likely to contribute to familial aggregation, including rare genetic variants (23–25), nonadditive effects, and shared environment.

Expanding GWAS, either through new genotyping or meta-analyses, should continue to yield new susceptibility loci. The pace of expected discovery, measured by the number of additional associations for given study sizes, will likely continue to be faster for those cancers with the highest FRRs. For stomach cancer (Asian population), osteosarcoma, testes, and CLL, all with FRRs over 2, we predict that more loci may be identified for studies of comparable size (15). Consistent with this prediction, comparatively smaller GWAS of these highly heritable cancers have historically yielded more associations.

Genetic variants can influence cancer risk, either indirectly through associations with nicotine dependence and smoking behavior, or directly through other mechanisms (26). We estimate that the genetic determinants of smoking behavior accounted for at least 24% (95% CI = 14% to 37%) and 7% (4% to 11%) of the total heritability of lung and bladder cancers, respectively, in our study populations. These percentages likely underestimate the true influence of smoking genetics by not accounting for other, unmeasured, smoking characteristics, and may be slightly biased by systematic differences in the distribution of smoking characteristics between the general population and our GWAS control sample selected according to experimental design (Supplementary Table 17, available online). Our result is consistent with the largest meta-analysis (27) of case-control studies in which the excess relative risk (FRR = 1.82, 95% CI = 1.64 to 2.05) for lung cancer was reduced by 36% when focused on nonsmokers (FRR = 1.52, 95% CI = 1.11 to 2.06).

Our analysis of coheritability indicated that, in general, most pairs of cancers studied here are unlikely to have strong genetic correlations. Although our study had low power to detect modest but possibly important correlations between specific pairs of cancers (28), the distribution of test statistics over all pairs of cancers deviated little from the null (Figure 1), suggesting that at most a small fraction have modest or high (eg,  $\rho > 0.3$ ) correlations (Supplementary Tables 18–19, available online). Four pairs of sites that show notable correlation in our study are bladder and lung, testes and kidney, DLBCL and CLL, and DLBCL and osteosarcoma. Pleiotropic analyses of polygenic risk scores offered additional evidence for overlapping genetic architectures between bladder and lung cancer and between DLBCL and CLL. While no family study has been of sufficient size to explore the connection between DLBCL and CLL or DLBCL and osteosarcoma, the Icelandic registry study (6) found that the relative risks for testicular cancer were elevated in relatives of kidney cancer case patients, as indicated by our results.

Analysis of the genetic correlation for the same cancer site but across distinct subtypes and populations can provide important insights. For instance, the strong genetic correlation between aggressive and nonaggressive prostate cancer indicates that common SNPs are unlikely to offer a diagnostic means to distinguish these two subtypes of cancer with different prognoses. On the other hand, there was an absence of correlation for lung cancer between studies in nonsmoking Asian females and a Caucasian population where case patients were primarily smokers. This result is consistent with the observation that distinct sets of susceptibility SNPs have emerged in these two populations (29,30). In general, low correlation across populations can be caused by different causal SNPs, differences in effect sizes or

allele frequencies of shared causal SNPs, and differences in linkage disequilibrium (LD). Simulations (results not shown) based on empirical genotype data suggest that the observed low correlation ( $\rho = 0.10$ , SE = 0.15) is unlikely to be explained by general differences in the allele frequency and LD pattern between these two populations. However, substantial confounding may arise if disease-causing loci are particularly selected to have strong differences with respect to be these factors.

Our study has some limitations. While total sample size was large, the number of cases for certain cancer sites, especially when stratified by subtypes, could be limited, resulting in large confidence intervals for estimates of heritability. Such uncertainty, together with various sources of bias and variability of the prior estimates of FRR from registry-based studies, make it more difficult to assess the proportion of observed FRR that can be attributed to common SNPs. Because of the limited sample size, our analysis also lacked power to detect modest genetic correlations among specific cancer types/subtypes. This lack of precision may partially explain the more extensive familial coaggregation observed in a study based on the Icelandic Cancer Registry (6). However, the discrepancy also can be attributed to different cancers pathologies, environmental differences across populations, and, most importantly, other sources of shared heritability. Another limitation is that principal components may not perfectly adjust for ancestry and therefore shared environments or behaviors among individuals with similar ancestry may inflate our estimates of heritability. However, tests for such inflation found little evidence to support this concern.

For some cancer sites, the individuals used in the previous study (13) of GWAS heritability overlap with a subset the individuals used here, and therefore the two studies do not offer independent assessments of heritability. For kidney cancer, our population overlaps considerably with that studied by Lu and colleagues (13), who estimated its heritability to be 0.18. For other cancers, including bladder, lung, pancreas, and prostate, the prior study populations are primarily subsets of our own populations, including the 25% to 50% of individuals from the earliest GWAS. Prior estimates for these four cancers were respectively 0.01, 0.10, 0.18, and 0.81, and show substantive differences from our updated estimates (Table 1). This analysis of more than 80 000 individuals provides important perspectives on the heritability of cancer across anatomical sites. We affirm that there is a large heritable component to most cancers, but the majority of cancer heritability cannot be attributed to known susceptibility loci. We further demonstrate that marker SNPs are not omnipresent across cancers, in that there does not appear to be strong genetic correlations between most pairs of cancer sites. Among smoking-related cancers, we showed that the genetic determinants of smoking make a statistically significant contribution to the heritability of lung and bladder cancer. Overall, as GWAS expand in size and design, a comprehensive analysis should continue to focus on cancers individually, look across groups of cancers identified as being genetically similar, and account for important environmental risk factors.

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