

Received: 28 August 2018 Accepted: 25 March 2019 Published online: 11 April 2019

OPEN Exome array analysis of rare and low frequency variants in amyotrophic lateral sclerosis

Annelot M. Dekker¹, Frank P. Diekstra¹, Sara L. Pulit₀¹, Gijs H. P. Tazelaar¹, Rick A. van der Spek¹, Wouter van Rheenen 1, Kristel R. van Eijk¹, Andrea Calvo 1, Maura Brunetti², Philip Van Damme^{3,4,5}, Wim Robberecht^{3,4,5}, Orla Hardiman^{6,7}, Russell McLaughlin⁸, Adriano Chiò 602, Michael Sendtner 609, Albert C. Ludolph 10, Jochen H. Weishaupt 10, Jesus S. Mora Pardina¹¹, Leonard H. van den Berg¹ & Jan H. Veldink¹

Amyotrophic lateral sclerosis (ALS) is a fatal neurodegenerative disease that affects 1 in ~350 individuals. Genetic association studies have established ALS as a multifactorial disease with heritability estimated at ~61%, and recent studies show a prominent role for rare variation in its genetic architecture. To identify rare variants associated with disease onset we performed exome array genotyping in 4.244 cases and 3.106 controls from European cohorts. In this largest exome-wide study of rare variants in ALS to date, we performed single-variant association testing, gene-based burden, and exome-wide individual set-unique burden (ISUB) testing to identify single or aggregated rare variation that modifies disease risk. In single-variant testing no variants reached exome-wide significance, likely due to limited statistical power. Gene-based burden testing of rare non-synonymous and loss-of-function variants showed NEK1 as the top associated gene. ISUB analysis did not show an increased exome-wide burden of deleterious variants in patients, possibly suggesting a more regionspecific role for rare variation. Complete summary statistics are released publicly. This study did not implicate new risk loci, emphasizing the immediate need for future large-scale collaborations in ALS that will expand available sample sizes, increase genome coverage, and improve our ability to detect rare variants associated to ALS.

Amyotrophic lateral sclerosis (ALS) is a rapidly progressing and fatal neurodegenerative disease with an estimated lifetime risk of approximately 1 in 3501. The central hallmark of the disease is dysfunction and gradual loss of upper and lower motor neurons, resulting in progressive muscle weakness and eventual death due to respiratory failure. Median survival is three to five years after symptom onset^{2,3}. A subset of patients shows clinical signs of extramotor system involvement, most notably cognitive impairment⁴. There is no cure to date, and the sole effective drug, riluzole, only extends survival by approximately three months⁵.

ALS is considered a complex disease to which both environmental and genetic risk factors contribute to disease susceptibility. Twin-based heritability is estimated at approximately 61%. Over the past decades multiple genes and genomic regions have been implicated in ALS via linkage studies, genome-wide association studies (GWAS), and more recently, large scale sequencing efforts⁷. A recent GWAS in ALS, in approximately 14,000 cases and 30,000 controls, identified a total of 7 loci associated to the disease. The top loci together explain 0.2%

¹Department of Neurology, Brain Center Rudolf Magnus, University Medical Center Utrecht, Utrecht, The Netherlands. ²Rita Levi Montalcini' Department of Neuroscience, ALS Centre, University of Torino, Turin, Italy. ³KU Leuven - University of Leuven, Department of Neurosciences, Experimental Neurology and Leuven, Research Institute for Neuroscience and Disease (LIND), B-3000, Leuven, Belgium. 4VIB, Vesalius Research Center, Laboratory of Neurobiology, Leuven, Belgium. ⁵University Hospitals Leuven, Department of Neurology, Leuven, Belgium. ⁶Academic Unit of Neurology, Trinity College Dublin, Trinity Biomedical Sciences Institute, Dublin, Ireland. ⁷Department of Neurology, Beaumont Hospital, Dublin, Ireland. ⁸Population Genetics Laboratory, Smurfit Institute of Genetics, Trinity College Dublin, Dublin, Ireland. 9Institute of Clinical Neurobiology, University of Würzburg, Würzburg, Germany. ¹⁰Department of Neurology, Ulm University, Ulm, Germany. ¹¹ALS Unit, Hospital San Rafael, Madrid, Spain. Leonard H. van den Berg and Jan H. Veldink jointly supervised this work. Correspondence and requests for materials should be addressed to J.H.V. (email: J.H.Veldink@umcutrecht.nl)

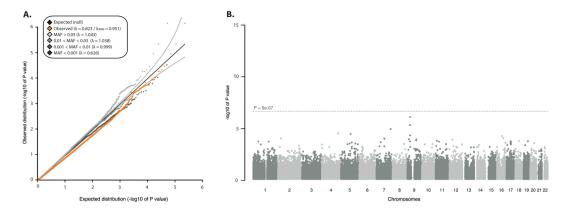


Figure 1. Quantile-quantile plot and Manhattan plot of p-values of the single-variant association analysis. (**A**) Quantile-quantile plot of single-variant association analysis using logistic regression binned by minor allele frequency. (**B**) Manhattan plot of p-values of exome-wide association testing comprising 7350 individuals (4,244 cases and 3,106 controls) and 100,896 non-monomorphic variants. The *x* axis depicts chromosomal position and the *y* axis shows the significance of association derived by logistic regression. The dotted line corresponds to the exome-wide significance threshold of $p = 5 \times 10^{-7}$).

of the estimated 8.5% SNP-based heritability; heritability estimates across all SNPs indicate that the bulk of the SNP-based heritability is captured in low-frequency (defined as minor allele frequency (MAF) 1-10%) variants beneath genome-wide significance. These findings are consistent with a disease architecture in which rare and low-frequency variants play an important role.

Imputation of low-frequency (MAF 0.5–5.0%) variants in genotyping data has become more accurate with the increasing public availability of large reference panels, and with the combination of public reference panels with study-specific reference panels 10,11 . High quality imputation of rare variants, (MAF < 0.5%), however, has remained challenging 12,13 . The exome array provides a low-cost alternative to large-scale exome or genome sequencing, albeit with lower resolution. In this largest exome-wide study in ALS to date, we used the exome array to investigate the role of low-frequency and rare variants by genotyping over 240,000 primarily functional coding variants in 7350 ALS patients and controls drawn from different European cohorts. Using this approach, we aimed to further elucidate the role of these low-frequency and rare variants in the genetic landscape of ALS.

Results

Analyses comprised 4,244 cases and 3,106 controls from 6 European countries. Of 242,901 genotyped sites, 233,331 sites passed quality control, of which 100,896 were non-monomorphic in cases and controls. A breakdown of the included samples per country is provided in Supplementary Table 1.

Single-variant testing. We performed single-variant testing using logistic regression, adjusting for sex and population principal components, and assuming an additive model. We found the genomic inflation factor ($\lambda_{\rm GC}$) to be 0.823 (Fig. 1A), reflecting the high prevalence of rare (i.e., MAF < 0.5%) variants in the data and therefore low power to detect effects at these rare alleles. No variants reached exome-wide significance (p < 5 × 10⁻⁷; Fig. 1B). Multiple top-associated signals (MAF 0.23–0.29) were located on chromosome 9 (lowest p-value 1.72 × 10⁻⁶; Supplementary Table 2). These signals represent the large intronic hexanucleotide repeat expansion in the gene *C9orf72*, the most common genetic cause of ALS identified to date ^{14,15}. The strongest associated rare variants (MAF < 0.5%), rs200161705 (odds ratio (OR) = 2.74, p = 5.76 × 10⁻⁵) and rs181906086(OR = 0.35, p = 1.68 × 10⁻⁴), are both missense variants as annotated by Variant Effect Predictor ¹⁶ and are located in the genes *NEK1* and *CAPN14* respectively (Supplementary Table 2).

Previous work has shown that, in testing rare variation down to minor allele count <400, the Firth test has the best combination of type I error and power in balanced and moderately unbalanced case-control studies¹⁷. We therefore repeated our analysis using Firth logistic regression and found the results to be highly consistent with our initial regression results (data not shown).

Supplementary Table 3 shows the results from one of the largest ALS GWAS to date (Van Rheenen *et al.*⁸) compared to the results from our single variant analysis⁸.

Gene-based analysis. Next, to test for an association between aggregated rare variants clustering in genes and ALS, we performed gene-based analysis using the unified optimal sequence kernel association test (SKAT-O)¹⁸. As SKAT-O is a combination of burden and variance statistics, the test retains discovery power under different genetic architecture models (i.e., by varying the proportion of causal variants within the gene that either have the same direction of effect or presence of both protective and deleterious variants, and then testing for association). We observed no evidence for residual population stratification based on the standardized genomic inflation factor λ_{GC} of 1.016 after correction for principal components, sex and cohort (Supplementary Figure 1). No individual gene surpassed the predefined Bonferroni-corrected significance threshold ($p = 3.45 \times 10^{-6}$, after adjusting for 14,488 genes). For the top genes most associated to ALS ($p < 1 \times 10^{-3}$; Table 1), we calculated exact p-values using 500,000 case-control permutations. The strongest signal in the gene-based burden test using

Gene	Chr	Start (bp)	End (bp)	Nominal p value SKAT-O	Exact p value SKAT-O	Nominal p value Firth test	Number of variants
NEK1	4	170314421	170533778	2.73×10^{-5}	1.21×10^{-5}	3.06×10^{-5}	6
ZNF287	17	16453626	16472520	3.08×10^{-5}	3.21×10^{-5}	2.06×10^{-5}	2
CAPN14	2	31395922	31456724	2.16×10^{-4}	1.73×10^{-4}	1.18×10^{-3}	5
IRF8	16	85932774	85956212	2.50×10^{-4}	2.25×10^{-4}	1.34×10^{-3}	4
F10	13	113777113	113803843	5.94×10^{-4}	4.77×10^{-4}	5.20×10^{-4}	4
PTPRB	12	70910630	71031220	7.10×10^{-4}	5.74×10^{-3}	0.37	21
SYNM	15	99645286	99675800	7.55×10^{-4}	6.14×10^{-3}	5.05×10^{-4}	8
SGK3	8	67624653	67774257	7.90×10^{-4}	7.00×10^{-3}	8.52×10^{-4}	2

Table 1. Top genes associated to ALS, as found through gene-based burden testing. Results of gene-based burden testing using SKAT-O and Firth test, results limited to genes exceeding $p < 1 \times 10^{-3}$ using SKAT-O. Positions given for human build 37. Exact p values generated from 500,000 case-control permutations.

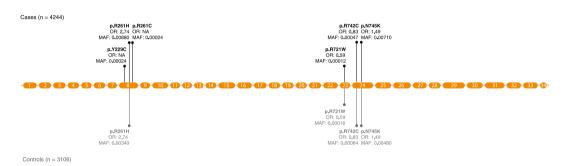


Figure 2. Gene mutation plot of *NEK1*. The orange circles represent exons relative to their size in basepairs. The vertical lines represent the distribution of SNVs across *NEK1* in the gene-based test (SKAT-O), with their corresponding amino acid change, minor allele frequency (MAF) and odds ratio (OR) in the single-variant analysis.

SKAT-O was NEK1 (p = 1.21×10^{-5} ; Fig. 2). For the majority of the top associating genes repeating the analysis using Firth test showed highly concordant results (Table 1). Conditioning the gene-based analysis on variants in that gene present as top hit in the single-variant analysis resulted in an expected rise in p-value, due to loss in power after conditioning on the variant with the highest MAF (Supplementary Table 4).

Resolution for gene-based analysis was assessed by comparing the number of missense and loss-of-function variants in previously identified ALS genes (i.e., SOD1, FUS, TARDBP, KIF5A, NEK1 and C21orf2) present in ExAC to the number of variants tested in our gene-based analysis (Supplementary Table 5). SKAT-O analysis in these genes comprised a maximum of 6 variants (NEK1; $p = 1.21 \times 10^{-5}$).

Individual set-unique burden analysis. Previous studies have indicated that in some complex diseases, disease-affected individuals carry an excess number of missense and nonsense variants¹⁹. Therefore, we next sought to investigate the exome-wide burden per individual of particular classes of variants by estimating the individual set-unique burden (ISUB) for each individual. We first annotated all variants available for analysis. Of the 24,844 variants carried only by the 4,244 ALS cases, 23,231 variants are nonsynonymous and 6,600 are deleterious, according to the CONDEL algorithm²⁰; these variants comprise the 'set-unique' variants in cases. Of the 11,856 set-unique variants in the 3,106 controls, 11,088 are nonsynonymous and 3,135 are deleterious. We used CONDEL to assign deleteriousness scores to each variant and then summed and normalized these scores. Finally, for all set-unique nonsynonymous and loss-of-function variants (called the 'NS' group of variants) as well as for the observed loss-of-function variants predicted to be deleterious (the 'DEL' group of variants), we calculated the exome-wide individual set-unique burden (ISUB) per individual.

We found the ISUB scores based on the NS group of variants to be significantly higher in ALS patients compared to controls ($p=9.6\times10^{-137}$). We observed no such difference for ISUB calculated from the DEL group of variants (p=0.21) (Table 2, Supplementary Figure 3). To control for the case-control imbalance in our sample (case-control ratio ~1.36:1) we repeated the ISUB analysis in balanced cohorts only (i.e., including 5,069 individuals from the Dutch, Belgian, and Irish cohorts with case-control ratios nearing 1:1). In this balanced subset, we observed no significant difference in ISUB score based either on nonsynonymous variants (p=0.37) or on deleterious variants (p=0.51, Table 2, Supplementary Figures 2 and 4).

We additionally confirmed that the differing number of observed variants per individual (N = 7 and N = 4 respectively, $p = 1.38^{-276}$) in the imbalanced (i.e., complete) dataset drove the significant difference in set-unique burden between cases and controls (Supplementary Table 6). Balancing the case-control ratio eliminated this difference (N = 7 for nonsynonymous variants in both the case and control datasets, p = 0.19) (Supplementary Table 6, Supplementary Figure 5). The scores of deleteriousness per variant, as measured by CONDEL, did not

		All Cohorts				Balanced Cohorts			
Variant type	Phenotype	No. of individuals	No. of SNVs	ISUB score (mean/median/sd)	p value	No. of individuals	No. of SNVs	ISUB score (mean/median/sd)	p value
DEL	Case	4244	6600	1.58/1.23/1.57	0.21	2489	4201	1.47/1.20/1.62	0.51
	Control	3106	3135	1.38/1.13/1.49	0.21	2580	4025	1.42/1.21/1.20	
NS	Case	4244	23,231	4.33/3.49/4.16	9.6×10^{-137}	2489	14,831	3.97/3.43/4.55	0.37
	Control	3106	11,088	2.30/1.89/2.10	J.U ∧ 10 ···	2580	14,310	3.83/3.49/2.82	

Table 2. Results individual set unique burden analysis. Results given for analysis comprising all individuals (all cohorts; N=7350) and for a subset of samples comprising balanced case-control cohorts only (balanced cohorts; samples from The Netherlands, Belgium and Ireland, N=5069). P values given for logistic regression with first six principal components and country of sample origin included as covariates. DEL = deleterious variants, NS= all nonsynonymous and loss-of-function variants.

differ between cases and controls for the imbalanced dataset (p = 0.58) and the balanced dataset (p = 0.87) (Supplementary Table 7, Supplementary Figure 5). Repeating the analysis after removing outlying scores (defined as >5 standard deviations from set mean ISUB score) or logarithmic transformation of the data did not change the results

Discussion

Using the exome array in 7,350 ALS cases and matched controls, we sought to investigate the role of low-frequency and rare variants in the etiology of ALS. Despite being the largest study of rare variants in ALS to date, no associations reached the predefined significance levels in single-variant association testing, gene-based burden testing and exome-wide individual set-unique burden (ISUB) testing. The strongest associating signals have previously been linked to ALS.

In accordance with other exome array studies, we had minimal power to find significant associations at rare, modest effect variants. At a minor allele frequency (MAF) of 1%, variants with large effect sizes (odds ratio >2.3) could be detected with 80% power; larger effect sizes are necessary to have sufficient power (>80%) for variants with lower frequency. Because the exome array content is nearly entirely comprised of rare variants that, by definition, typically have low LD with other variants, our power to detect other associated variants with ALS through LD-tagging was also low. Given our results, we can conclude that none of the rare variants captured on the exome array are large-effect variants associated with ALS.

For gene-based burden testing our analysis yielded sufficient power to detect genes with a large proportion of causal variants (\geq 50%) contributing to disease risk (including genes with \leq 25% protective variants). For genetic architectures with lower percentages of causal variants and higher percentages of protective variants, power was low. Different genes associated to ALS likely contain different genetic architectures. For example, previous work showed that mutations in several ALS-related genes predominantly affect specific regions of the gene (e.g. mutations causing exon 7 skipping in *FUS* and mutations primarily located at the C-terminal cargo-binding region in *KIF5A*), whereas in other genes (e.g. *SOD1*) mutations are more dispersed^{21–24}. This architectural complexity makes selecting appropriate settings for power analyses difficult; while we can evaluate power for simpler architectures (e.g., genes containing mostly causal risk variation), it is difficult to evaluate power for a gene with a much more complex architecture (e.g., with both risk and protective variants that must occur on specific haplotypes and/or in combination with specific genetic backgrounds)²⁵.

Additional limitations include the fact that of the ~240,000 variants present on the array; the majority were monomorphic in our dataset. Illustrative of this low resolution is the observation that for gene-based burden testing of *KIF5A*, a gene recently implicated in ALS pathogenesis, our dataset only included three variants. Two of these three variants are located outside the previously identified ALS-associated domain in the C-terminus of the protein^{23,24}. As a result, we did not find an association between *KIF5A* and ALS. Our analyses also only include European-ancestry samples; given the strong geographic localization of rare variants, such a design leaves the many rare variants found only in non-European samples untested. Further, since most of the rare variants that have been identified in rare variant association studies to date have modest-to-weak effect sizes, much larger sample sizes or denser coverage of the genome are necessary in order to robustly implicate novel associations that increase risk of disease.

Despite power limitations, we can still draw several key conclusions. First, we confirm several of the strongest associations previously found in ALS, demonstrating the validity of our approach. The common variants associating most strongly with ALS in this study map to the gene *C9orf72*, a well-established genetic risk factor in ALS. These signals are driven by a massive hexanucleotide repeat (GGGGCC) expansion between non-coding exons 1a and 1b, an expansion which is the most frequent genetic cause of ALS discovered to date²⁶. The strongest associated rare variant (rs200161705) located in *NEK1* was recently discovered as a risk variant in an inbred population in The Netherlands and subsequently replicated in an international cohort including ALS patients with and without a positive family history²⁷. Since these initial findings, several studies have confirmed *NEK1* as an ALS risk gene^{27–29}. In our study, *NEK1* was the top hit in the gene-based analysis, demonstrating how, in some instances, application of the exome array can be a more efficient approach than sequencing for finding large effect variants. Future studies that use arrays with both exome and common variant content, to allow for not only interrogation of rare coding variants but also interrogation of common variation through genotyping and imputation, will be well positioned to identify signals like that residing in *NEK1*.

Because CAPN14 gene was among the top findings in both single-variant and gene-based analysis, it represents a tantalizing finding but one that requires additional follow-up in larger samples. CAPN14 is a member of the calpain family; calpains are cysteine proteases that are activated by calcium, possibly through increased glutamatergic neuronal transmission. Calpains have been identified as regulators of axonal survival in injury-induced and developmental degeneration via necrotic and apoptotic pathways^{30–32} and an ALS mouse model has shown calpain inhibition to be neuroprotective³³. Calpain14 is primarily expressed in esophageal mucosa and has also been associated with eosinophilic esophagitis^{34,35}. Although an association between the calpain family and ALS is possible, future studies will be necessary to further establish and later unravel a potential role for the gene in disease etiology. While our ISUB analysis did not reveal an exome-wide increase in individual set-unique burden between cases compared to controls, our results do demonstrate the importance of careful study design, in particular of case-control studies seeking to investigate rare variation. A case-control imbalance, allowing for the detection of more rare variants in one group than the other, can induce spurious associations if not handled properly. An identical analysis of genome-wide individual burden of nonsynonymous and loss-of-function variants performed in schizophrenia found a difference between cases and controls in a comparatively smaller sample $(N = 2,003)^{19}$. Despite the described genetic overlap of ALS and schizophrenia, the genetic architecture of schizophrenia is likely quite different from ALS³⁶. GWAS in schizophrenia have revealed a prevalent role for common variants in the disease, and schizophrenia is also more heritable (h²~80%)³⁷. Given the highly polygenic architecture of schizophrenia, an ISUB analysis in that phenotype may better capture the many disease-associated loci scattered across the genome. In contrast, the architecture of ALS may be less polygenic and/or more region-specific, and inclusion of exome-wide variation may contribute additional noise to the analysis, thus obscuring potential statistical signal. Additional studies that help to first pinpoint these potential localized signals are necessary to test this hypothesis.

With the lowering costs of high-coverage sequencing and the increasing availability of large reference panels that allow for imputation of tens of millions of low-frequency variants, high-quality studies of rare variants in ALS are expected in the coming years³⁸. These future larger and more resolute studies are expected to allow additional data filtering in order to separate noise from true signal and will be better powered to accommodate the diverse genic architectures observed in ALS to date. Given that our analysis of functional coding variants in thousands of European ALS patients and controls did not implicate new risk loci, our analyses reaffirm the need for future studies of rare variants in ALS that are internationally collaborative, allowing for the collection of larger and more ancestrally diverse sample sizes and denser coverage of the genome. In order to facilitate large scale rare variant analyses summary statistics will be made publicly available through the Project MinE databrowser website at http://databrowser.projectmine.com/.

Methods

Study population. All 4495 patients and 3227 controls included in this study were recruited at specialized neuromuscular centers in Belgium, Germany, Ireland, Italy, Spain and the Netherlands. Patients were diagnosed with possible, probable or definite ALS according to the 1994 El-Escorial criteria³⁹. All controls were free of neuromuscular diseases and matched for age, sex and geographical location.

Genotyping and quality control. We conducted genotyping using Illumina HumanExome-12v1 BeadChips in accordance with the manufacturer's protocol. A description of the exome chip design is available from http://genome.sph.umich.edu/wiki/Exome Chip Design. We applied the GenTrain 2.0 clustering algorithm for genotype calling as implemented in the Illumina GenomeStudio software package. Initial genotype calls were made based on the HumanExome cluster file provided by Illumina. More accurate cluster boundaries were determined based on the actual study data by exclusion of samples with a GenCall quality score in the lowest 10th percentile of genotyped variants (p10GC < 0.38) or call rate < 0.99. The final genotype calls of the whole dataset were obtained using these more precise cluster boundaries. We then performed additional sample and variant quality control (QC) was performed using PLINK 1.9, excluding all samples with missing call rates higher than 5% or sex discordance^{40,41}. A subset of independent, high quality SNPs (non-AT/CG, autosomal variants not located in high linkage disequilibrium (LD) regions with minor allele frequency (MAF) > 0.05, genotyping rates >99%, R^2 < 0.05 and Hardy-Weinberg equilibrium p values > 1 × 10⁻³) was used to determine heterozygosity rates, population stratification, intersample relatedness and sample duplication. Closely related and duplicated samples (pi-hat > 0.2) and samples failing heterozygosity checks (>4 standard deviations) were removed. To assess population stratification, we calculated principal components on all remaining individuals and HapMap3 individuals using EIGENSTRAT^{42,43}. We removed population outliers (defined as >10 standard deviations from the HapMap CEU mean on PC 1-4) (Supplementary Figure 8a). Repeating principal component analysis with data from The 1000 Genome Project did not result in the removal of additional population outliers (Supplementary Figure 8b). Non-autosomal variants and variants with haploid heterozygous calls, call rate <98%, deviation from Hardy-Weinberg equilibrium in controls (p $< 1 \times 10^{-6}$) and biased missingness between cases and controls (p $< 5 \times 10^{-3}$) were removed from the dataset. The final dataset comprised a total of 7,350 individuals (4,244 cases and 3,106 controls) and 233,331 variants.

Statistical analyses. We performed single-variant association testing in PLINK 1.9 using logistic regression under an additive genetic model 40,41 . Covariates were selected based on significant correlation to the phenotype using logistic regression. To correct optimally for population stratification and batch effects due to sample handling, we used the first five principal components, sex and country of sample origin as covariates. We tested all non-monomorphic variants in both cases and controls (N = 100,896). We corrected for multiple testing using a Bonferroni adjusted p value threshold (p = $0.05/100,896 = 4.95 \times 10^{-7}$), which nears the exome-wide significance threshold of 5×10^{-7} . Genomic inflation factors were calculated using R v3.2.2 (http://www.r-project.

org). Since previous work showed that Firth test has the best combination of type I error and power in balanced and moderately unbalanced studies of rare variant (i.e. variants with a minor allele count <400) we repeated the single-variant analysis using Firth logistic regression in R v3.4.1. as implemented in the logistf package.

Gene-based analysis was conducted using the unified optimal sequence kernel association test (SKAT-O) as implemented in the R-package $SKAT^{44}$. Variants were functionally annotated using Ensembl Variant Effect Predictor (VEP)¹⁶. All non-synonymous and loss-of-function variants with a MAF < 0.01 and minor allele count >1 were included. Genes containing one annotated variant were excluded. The model was run with sex, the first five principal components and country included in the model as covariates. Exact p-values were generated from 500,000 case/control permutations for hits with a nominal p-value < 0.001. Results were considered significant after Bonferroni correction for number of tests (p = 0.05/14,488 = 3.45 × 10⁻⁶). Burden testing was repeated using Firth logistic regression in R v3.4.1.

To investigate the exome wide burden of rare single nucleotide variants (SNVs) in ALS cases compared to controls an individual set-unique burden (ISUB) analysis was performed as described previously¹⁹. Only variants uniquely present in controls or uniquely present in cases were selected for this analysis. Allele frequencies were compared to the non-Finnish European cohort in ExAC and variants with a MAF higher than 0.5% were removed⁴⁵. Deleteriousness of the variants was assessed using the CONDEL algorithm, which assigns a weighted average of normalized scores from multiple prediction tools to each coding variant, as well as a 'neutral' or 'deleterious' label²⁰. Non-scored loss-of-function variants were assigned a deleterious label and the maximum score occurring in the corresponding dataset. The final individual burden score was computed by summing the scores for all observed non-synonymous and loss-of-function set-unique variants (NS) as well as for the set-unique variants predicted to be deleterious (DEL).

We assigned scores under a dominant model, assigning a score only once in case of a homozygous genotype. The analysis was first performed in the complete dataset comprising 4,244 cases and 3,106 controls. Since the case-control ratio is an important determinant in whether a variant appears to be unique, we repeated all analyses in a balanced case-control subset of Dutch, Belgian and Irish samples (2,489 cases and 2,580 controls). Differences in ISUB scores between cases and controls were assessed using logistic regression with the first five principal components and country included in the model as covariates. To reduce the effect of outlying scores on the results the analyses were repeated with outlying scores (>5 standard deviations from set mean) removed as well as with logarithmic transformation of the data. R version 3.2.2 was used for all statistical analyses (http://www.r-project.org).

Power analysis. To estimate detectable effect sizes given the low frequencies of rare single nucleotide variants, (Supplementary Fig. 6), we performed power analysis for single-variant associating testing using the online Genetic Power Calculator (http://zzz.bwh.harvard.edu/gpc/)⁴⁶.

Power analysis for gene-based burden analysis using SKAT-O as implemented in the R-package SKAT revealed sufficient power (>86% at 200 simulations) to detect genes with a large proportion (\geq 50%) of causal (i.e., risk-increasing) variants among rare variants (defined as MAF < 0.01) contributing to disease risk, including genes with \leq 25% protective (i.e., risk-decreasing) variants (Supplementary Figure 7) (see package documentation for information regarding assumptions, default settings and methods at https://CRAN.R-project.org/package=SKAT).

Ethics statement. The study was approved by the Medical Ethical Committee from the University Medical Center Utrecht, The Netherlands. Also, the present study followed study protocols approved by Medical Ethical Committees for each of the participating institutions. Written informed consent was obtained from all participating individuals. All methods were performed in accordance with the relevant national and international guidelines and regulations.

Data Availability

Summary statistics for this study are publicly available through the Project MinE Data Browser at http://data-browser.projectmine.com/.

References

- 1. Johnston, C. A. et al. Amyotrophic lateral sclerosis in an urban setting. J Neurol 253, 1642–1643 (2006).
- 2. Huisman, M. H. B. et al. Population based epidemiology of amyotrophic lateral sclerosis using capture-recapture methodology. Journal of Neurology, Neurosurgery & Psychiatry 82, 1165–1170 (2011).
- 3. Rooney, J. et al. Survival Analysis of Irish Amyotrophic Lateral Sclerosis Patients Diagnosed from 1995–2010. PLoS ONE 8, e74733–10 (2013).
- 4. Abrahams, S., Newton, J., Niven, E., Foley, J. & Bak, T. H. Screening for cognition and behaviour changes in ALS. Amyotrophic Lateral Sclerosis and Frontotemporal Degeneration 15, 9–14 (2014).
- Miller RG, M. J. M. D. Riluzole for amyotrophic lateral sclerosis (ALS)/motor neurondisease (MND). The Cochrane Library 1–36 (2012).
- Al-Chalabi, A. et al. An estimate of amyotrophic lateral sclerosis heritability using twin data. Journal of Neurology, Neurosurgery & Psychiatry 81, 1324–1326 (2010).
- Al-Chalabi, A., van den Berg MD, P. L. H. & Veldink, J. Gene discovery in amyotrophic lateral sclerosis: implications for clinical management. *Nature Reviews Neurology* 1–9, https://doi.org/10.1038/nrneurol.2016.182 (2016).
- 8. van Rheenen, W. et al. Genome-wide association analyses identify new risk variants and the genetic architecture of amyotrophic lateral sclerosis. Nat Genet 1–8, https://doi.org/10.1038/ng.3622 (2016).
- 9. Al-Chalabi, A., Calvo, A., Chio, A., Colville, S. & Ellis, C. M. Analysis of amyotrophic lateral sclerosis as a multistep process: a population-based modelling study. *The Lancet* 13, 1108–1113 (2014).
- 10. McCarthy, S. et al. A reference panel of 64,976 haplotypes for genotype imputation. Nat Genet 1–7, https://doi.org/10.1038/ng.3643 (2016).

- 11. Huang, J. et al. Improved imputation of low-frequency and rare variants using the UK10K haplotype reference panel. Nature. Communications 6, 1–9 (2015).
- 12. Auer, P. L. Rare variant association studies: considerations, challenges and opportunities. Genome Med 7, 1-11 (2015).
- Hoffmann, T. J. & Witte, J. S. Strategies for Imputing and Analyzing Rare Variants in Association Studies. Trends in Genetics 31, 556–563 (2015).
- 14. Renton, A. E. et al. A Hexanucleotide Repeat Expansion in C9ORF72 Is the Cause of Chromosome 9p21-Linked ALS-FTD. Neuron 72, 257–268 (2011).
- 15. DeJesus-Hernandez, M. et al. Expanded GGGGCC Hexanucleotide Repeat in Noncoding Region of C9ORF72 Causes Chromosome 9p-Linked FTD and ALS. Neuron 72, 245–256 (2011).
- 16. McLaren, W. et al. The Ensembl Variant Effect Predictor. Genome Biol. 17, 1-14 (2016).
- Ma, C., Blackwell, T., Boehnke, M. & Scott, L. J. The GoT2D investigators. Recommended Joint and Meta-Analysis Strategies for Case-Control Association Testing of Single Low-Count Variants. Genet. Epidemiol. 37, 539–550 (2013).
- 18. Lee, S., Wu, M. C. & Lin, X. Optimal tests for rare variant effects in sequencing association studies. Biostatistics 13, 762-775 (2012).
- 19. Loohuis, L. M. O. *et al.* Genome-wide burden of deleterious coding variants increased in schizophrenia. *Nature Communications* 6, 1–6 (2015).
- 20. González-Pérez, A. & López-Bigas, N. Improving the assessment of the outcome of nonsynonymous SNVs with a consensus deleteriousness score, Condel. Am. J. Hum. Genet. 88, 440–449 (2011).
- 21. Kaur, S. J., McKeown, S. R. & Rashid, S. Mutant SOD1 mediated pathogenesis of Amyotrophic Lateral Sclerosis. *Gene* 577, 109–118 (2016).
- 22. Zhou, Y., Liu, S., Liu, G., Öztürk, A. & Hicks, G. G. ALS-Associated FUS Mutations Result in Compromised FUS Alternative Splicing and Autoregulation. *PLoS Genet* 9, e1003895–17 (2013).
- 23. Nicolas, A. et al. Genome-wide Analyses Identify KIF5A as a Novel ALS Gene. Neuron 97, 1268-1282.e6 (2018).
- 24. Brenner, D. et al. Hot-spot KIF5A mutations cause familial ALS. Brain 141, 688-697 (2018).
- 25. Sham, P. C. & Purcell, S. M. Statistical power and significance testing in large-scale genetic studies. *Nature Publishing Group* 15, 335–346 (2014).
- 26. van Es, M. A. et al. Amyotrophic lateral sclerosis. The Lancet 0, (2017).
- 27. Kenna, K. P. et al. NEK1 variants confer susceptibility to amyotrophic lateral sclerosis. Nat Genet 1–8, https://doi.org/10.1038/ng.3626 (2016).
- Cirulli, E. T. et al. Exome sequencing in amyotrophic lateral sclerosis identifies risk genes and pathways. Science 347, 1436–1441 (2015).
- 29. Brenner, D. et al. NEK1mutations in familial amyotrophic lateral sclerosis. Brain 139, e28-e28 (2016).
- 30. Yang, J. et al. Regulation of Axon Degeneration after İnjury and in Development by the Endogenous Calpain Inhibitor Calpastatin. Neuron 80, 1175–1189 (2013).
- 31. Vosler, P. S., Brennan, C. S. & Chen, J. Calpain-Mediated Signaling Mechanisms in Neuronal Injury and Neurodegeneration. *Mol Neurobiol* 38, 78–100 (2008).
- 32. Wright, A. L. & Vissel, B. CAST your vote: is calpain inhibition the answer to ALS? J. Neurochem. 137, 140-141 (2016).
- 33. Rao, M. V., Campbell, J., Palaniappan, A., Kumar, A. & Nixon, R. A. Calpastatin inhibits motor neuron death and increases survival of hSOD1 G93Amice. *J. Neurochem.* 137, 253–265 (2016).
- 34. Lonsdale, J. et al. The Genotype-Tissue Expression (GTEx) project. Nature Publishing Group 45, 580-585 (2013).
- 35. Sleiman, P. M. A. et al. GWAS identifies four novel eosinophilic esophagitis loci. Nature Communications 5, 5593 (2014).
- 36. McLaughlin, R. L. *et al.* Genetic correlation between amyotrophic lateral sclerosis and schizophrenia. *Nature Communications* 8, 14774 (2017).
- 37. Ripke, S. et al. Biological insights from 108 schizophrenia-associated genetic loci. Nature 511, 421-427 (2014).
- 38. van Rheenen, W. et al. Project MinE: study design and pilot analyses of a large-scale whole- genome sequencing study in amyotrophic lateral sclerosis. European Journal of Human Genetics 7, 1–10 (2018).
- 39. Brooks, B. R. El Escorial World Federation of Neurology criteria for the diagnosis of amyotrophic lateral sclerosis. Subcommittee on Motor Neuron Diseases/Amyotrophic Lateral Sclerosis of the World Federation of Neurology Research Group on Neuromuscular Diseases and the El Escorial 'Clinical limits of amyotrophic lateral sclerosis' workshop contributors. *Journal of the Neurological Sciences* 124(Suppl), 96–107 (1994).
- Purcell, S. et al. PLINK: A Tool Set for Whole-Genome Association and Population-Based Linkage Analyses. The American Journal of Human Genetics 81, 559–575 (2007).
- 41. Chang, C. C. et al. Second-generation PLINK: rising to the challenge of larger and richer datasets. GigaSci 4, 7-16 (2015).
- 42. Price, A. L. et al. Principal components analysis corrects for stratification in genome-wide association studies. *Nat Genet* 38, 904–909 (2006).
- 43. Consortium, T. I. H. et al. Integrating common and rare genetic variation in diverse human populations. Nature 467, 52-58 (2010).
- 44. Lee, S. et al. Optimal unified approach for rare-variant association testing with application to small-sample case-control whole-exome sequencing studies. Am. J. Hum. Genet. 91, 224–237 (2012).
- 45. Lek, M. *et al.* Analysis of protein-coding genetic variation in 60,706 humans. *Nature* **536**, 285–291 (2016).
- Purcell, S., Cherny, S. S. & Sham, P. C. Genetic Power Calculator: design of linkage and association genetic mapping studies of complex traits. *Bioinformatics* 19, 149–150 (2003).

Acknowledgements

This project was supported by the Netherlands Organisation for Health Research and Development (Vici scheme to L.H. van den Berg). O. Hardiman is funded by the Health Research Board Clinician Scientist Programme and Science Foundation Ireland. The Thierry Latran Foundation (ALSIBD) and the ALS Association (2284) support R.L. McLaughlin. This project has received funding from the European Research Council (ERC) under the European Union's Horizon 2020 research and innovation programme (Grant Agreement No. 772376 - ESCORIAL. The authors would like to thank the Exome Aggregation Consortium and the groups that provided exome variant data for comparison. A full list of contributing groups can be found at http://exac.broadinstitute.org/about.

Author Contributions

F.D., W.R., P.D., R.M., O.H., R.M., A.C., M.S., J.W., J.M., J.V. and L.H. were involved in sample collection, phenotyping and data preparation. A.D., F.D., W.R., G.T., R.S., K.E. and S.P. performed data analyses. A.D. wrote the manuscript file with contributions from all authors. A.D., S.P., W.R. and R.S. prepared the figures. All authors read, reviewed and approved the final manuscript.

Additional Information

Supplementary information accompanies this paper at https://doi.org/10.1038/s41598-019-42091-3.

Competing Interests: L.H. van den Berg serves on scientific advisory boards for the Prinses Beatrix Spierfonds, Thierry Latran Foundation, Biogen and Cytokinetics; and serves on the editorial board of Amyotrophic Lateral Sclerosis And Frontotemporal Degeneration and The Journal of Neurology, Neurosurgery, and Psychiatry. J.H. Veldink reports that his institute received consultancy fees from Vertex Pharmaceuticals outside the submitted work. O. Hardiman has received speaking honoraria from Novarits, Biogen Idec, Sanofi Aventis and Merck-Serono, has been a member of advisory panels for Biogen Idec, Allergen, Ono Pharmaceuticals, Novartis, Cytokinetics and Sanofi Aventis and serves as Editor-in-Chief of Amyotrophic Lateral Sclerosis and Frontotemporal Dementia. Other authors have no reported conflicts of interest.

Publisher's note: Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.

Open Access This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons license, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons license, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons license and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this license, visit https://creativecommons.org/licenses/by/4.0/.

© The Author(s) 2019